Price and Quality Relationships in Local Service Industries

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Abstract: This working paper presents the findings of research on the relationship between price and quality in consumer service industries in the Washington, D.C. area. The study relies primarily upon consumer ratings of service provider quality and other data published in <u>Washington Consumer's Checkbook Magazine</u>. The data base includes nineteen service industries and, in virtually all cases, time series information for price and quality ratings over several ratings periods since the magazine's inception in 1976. The results provide interesting and frequently surprising information on basic price-quality relationships in this sector, and on the reliability of several non-price "signals" that consumers might use to gauge a service provider's probable performance. Specifically, only three of the nineteen industries report consistently significant positive price-quality correlations, and industries specializing in repair services frequently display significant negative correlations between price and quality. Further, possible signaling mechanisms, such as the size of a firm's Yellow Pages advertisement, or a firm's status as a member of a nationwide chain, do not function as indicators of higher quality in this data set.

JEL Classification: C10, D83, L15

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I. INTRODUCTION

In highly competitive markets characterized by low consumer search costs, we would expect product price and product quality to be strongly and positively correlated. If consumers can search easily among competitors and readily assess price and quality, those sellers providing higher quality will be able to charge a higher price to willing customers. Those failing to provide quality commensurate with price will lose customers and be forced to lower their price or exit the market.

If, however, consumers must incur significant costs to obtain accurate price and/or quality information, a high correlation between these two product attributes need not exist in equilibrium.¹ Presumably, the higher are search costs, the less likely consumers will be able to recognize good buys and bad buys before purchase, and the lower will be the correlation between price and quality. In the extreme case where consumers have no ability whatsoever to discern variations in quality pre-purchase, and assume simply that all products are of "average" quality, a variant of Gresham's Law will in sequential fashion drive higher quality providers from the market, since they will not be able to charge a sufficient premium to cover their greater production costs. Average quality will continue to decline until the products left in the market are of equally low quality. Although there would be a perfect correlation between price and quality at this degenerative point, this perfect correlation would signify a complete market failure [1].

A. Prior Research on Price-Quality Correlations

The degree of correlation between product price and quality in U.S. product markets has been explored extensively in a literature that dates back to 1950 [12]. The basic methodology of these studies is remarkably uniform. Product quality ratings usually are based on tests conducted by <u>Consumer Reports</u> ("CR").² The authors then calculate a correlation coefficient for these quality rankings and the list prices or, when available, the actual transaction prices for the tested

¹ The first theoretical treatment of price dispersion as an indicator of market imperfection appears in Stigler [16]. Stigler treated the initial existence of price dispersion as exogenous, and employed a simple consumer search model to explain why this condition would persist. Later treatments have attempted to explain the initial variation in prices. Usually, it is assumed that consumers differ in the amount of search they are willing to undertake, and that sellers specialize in serving either high or low-search consumers. Some sellers might, for example, bid for a particularly convenient location and charge a higher price to consumers who do not wish to search for a lower price. Other sellers with less advantageous sites will specialize in charging lower prices to consumers who are willing to search extensively to find the best deal. Entry will equalize the marginal return to the different strategies, and price dispersion will persist in equilibrium [4].

² A few of these studies use results from <u>Consumers Research</u>, and one article employed rankings of running shoes published in <u>Runners World</u> [2].

products. The results of these studies have also been quite consistent. In general, price and quality are only weakly correlated, with coefficients usually ranging from .20 to .25 [7]. These results are construed as evidence that consumer product markets perform "poorly," with one set of coauthors feeling motivated to title their article "The Chaos of Competition." [10]

This literature suffers from several weaknesses and limitations. Most obviously, the results can be no more reliable than CR's product ratings. Even if it is assumed that all of the CR protocols were appropriate and the tests performed competently, CR might not weight the various performance attributes for complex products such as stereo equipment, dishwashers, or tires in the same manner as would the average consumer.

And, if CR did happen to replicate the tastes of average consumers, significant subsets of buyers might still rank the products differently and willingly pay more for specialized products that might do relatively poorly in the ratings. Further, CR's ratings are not based on aesthetic qualities that may be very important to many consumers and that tend to be directly related to price.

Finally, sole reliance on CR automatically excludes all of the consumer service and retail industries that are supplied locally. Published studies of price and quality correlations for local services are limited to analyses of the legal and optometrical markets, where price advertising has been severely constrained [3]. We therefore have no systematic examinations of price and quality relationships in a major sector of the U.S. economy.

B. Current Research

The research reported in this working paper attempts to expand the limited scope and remedy some of the methodological weaknesses of prior studies of price-quality relationships in consumer industries. This analysis is made possible by the availability of an extensive data source for price and quality information for a large number of consumer service industries in the Washington, D.C. area. Such information has been published regularly by <u>Washington</u> <u>Consumers' Checkbook</u> since 1976.

<u>Washington Consumers' Checkbook</u> ("WCC") relies on quality evaluations supplied in response to mail questionnaires sent to all subscribers of <u>Consumer Reports</u> or WCC in the Washington area. Respondents are asked to rate service providers they have patronized during the last year as "superior," "adequate" or "inferior" on a variety of performance dimensions, such as "doing service properly," "starting and completing service promptly," and "overall performance." WCC "check rates" the firms that receive particularly high overall satisfaction ratings. Also reported is the number of consumers who rated each service provider.

The quality ratings are almost always accompanied by price information, usually in the form of an index. WCC staff contact each of the rated providers and ask for quotes on three or four standardized jobs. The providers' hourly labor rates are also listed where appropriate.

Although some establishments refuse to participate in this phase of WCC's research, ratings generally are available for 85-90 percent of the rated firms. WCC confers a check rating for price to the firms with the lowest price index scores.

In addition to price data, WCC customarily contacts local consumer protection offices to determine the number of complaints on file for each of the surveyed firms. (As discussed below, the complaint data provide one means of checking the reliability of the WCC survey results.) Finally, for approximately one-third of the rated industries, WCC gathers information from firms on the number of principal employees, such as auto mechanics or licensed plumbers. This information is used to normalize the complaint data for size of firm, and is not published separately.³ WCC agreed, however, to provide all available employee data for the most recent rating period. These data serve as a measure of firm size in the statistical analysis.⁴

WCC Magazine is published on a roughly biannual basis, with approximately five or six industries per issue. For certain industries, such as medical professionals and financial institutions, WCC does not provide overall quality ratings and/or price index information. For the current research, suitably complete data were available for 19 service industries. In most cases, each industry was rated on several occasions over time. These industries are presented in Table I below, with the number of individual data sets for each industry given in parentheses.

³ Assuming a firm has at least one complaint on file, the employee data can be derived using the absolute number of complaints and the complaint rate index, both of which are provided in the ratings. The majority of firms, however, have no complaints, thus precluding any indirect calculation of employee size.

⁴ The ratings already contain a very rough indicator of firm size as measured by the number of respondents rating the various providers. These numbers will, however, be heavily skewed toward providers who have received high ratings in past issues of WCC, since respondents to the WCC questionnaire will be far more likely than the general public to have relied on those prior ratings as a shopping tool.

TABLE ILocal Service Industries in Sample

- 1. Air Conditioning-Heating Repair (5)
- 2. Auto Body Repair (5)
- 3. Auto Mechanical Repair (8)
- 4. Major Appliance Repair (4)
- 5. Carpet Cleaning (3)
- 6. Carpet Sales, Installation (2)
- 7. Computer Repair (3)
- 8. Drycleaning (2)
- 9. Electricians (2)
- 10. Local Movers (3)

- 11. Locksmiths (3)
- 12. Pest Control (4)
- 13. Plumbers (4)
- 14. Restaurants (2)
- 15. Shoe Repair (2)
- 16. Supermarkets (3)
- 17. Tree Surgeons (1)
- 18. TV Repair (1)
- 19. Watch Repair (3)

WCC's approach to quality rankings avoids the major shortcoming of the <u>Consumer</u> <u>Reports</u> methodology. Since WCC quality ratings are based on actual consumer experience, they automatically reflect the weights that consumers assign to the various dimensions of firm performance. Although in some cases consumers may not be able to evaluate certain aspects of a service provider's competence and honesty as accurately as a trained tester, those ratings might still capture marked *differences* in quality among providers.

There is, however, a potentially high price to be paid for the consumer experience approach to rating service providers. The WCC "quality" ratings are really consumer "satisfaction" ratings. The ratings essentially measure how happy a consumer is after the service is completed and the bill is received. The difficulty is that the size of the bill may well influence the level of consumer satisfaction. Post purchase, most consumers will form a better opinion of, say, an auto repair establishment that presents a bill for \$150 than with one that charges \$450, even if there is no real difference in service quality.

This potential for bias exists irrespective of whether consumers have any clear knowledge of what competitors would have charged for the same job (although the magnitude of the bias would certainly be greater if patrons of the \$450 shop were aware that they had paid a \$300 premium). Higher bills inflict more pain, and the unfortunate recipients are less likely to reward providers with high ratings. Thus, depending on the strength of the bias, the WCC ratings could in theory reveal an inverse relationship between price and satisfaction even when the underlying relationship between price and true quality was positive.

Pre-purchase selection bias problems may also affect the satisfaction ratings. Consumers who choose to patronize higher cost establishments may also have higher expectations (and perhaps more complex service problems), and may therefore judge performance more critically, or be more likely to experience difficulties because of the nature of their particular service requests. Again, this potential bias in the WCC data may weaken or even reverse any positive relationship between price and quality that may actually exist in the surveyed service markets.

The price data published by WCC have their own potential problems. It would not be practical or prudent for WCC to gather price information as part of its consumer satisfaction survey. Even if consumers could remember price information reliably, there would in most instances be no way to pin down the precise nature and scope of the service that was provided. Thus, WCC gathers its own information, and is usually limited by time and budget constraints to soliciting price quotes for a small subset of the wide range of jobs provided by firms in most service industries. As a result, the price index based on these quotes is reliable only to the extent that the sampled prices reflect the prices of the services that consumers actually choose from each provider. As indicated earlier, WCC often publishes hourly labor data to supplement the price index, which provides a rough consistency check for the analysis.

The above difficulties not withstanding, the WCC ratings provide the most comprehensive source of information on price and quality in the local service sector that we are likely to have in the foreseeable future. It would not be economically feasible for a private tester such as Consumers Union to purchase such services directly and evaluate quality. There are simply too many providers and the cost of many of the services is too high. Further, unlike many of the durables tested by CR, services cannot be resold in the used market. The WCC data therefore deserve attention by researchers.

II. HYPOTHESES TESTED

A truly rigorous exploration of this subject area would begin with a complete structural model that would predict the equilibrium correlation between price and quality for firms in a given service industry based on production cost variables, consumer demand functions, and the magnitude of consumer search costs. Such a construct would allow us to test specific hypotheses concerning the absolute value of the correlation coefficient that we should expect in each of the the sample industries.

This study does not provide such a sophisticated underpinning, both due to the heroic scale of the theoretical challenges, and because suitable data would almost certainly not be available for many of the supply and demand variables that would be involved. Absent this theoretical construct, there is really very little we can say about the absolute correlations we would expect to see in these industries.

Presumably, none of the service industries included in this study fits the perfectly competitive model, if only because consumers must incur some search costs to determine firm

price and quality. And, in any event, we would not expect even a highly efficient market to function perfectly at any and every moment in time. Snapshot measurements such as those relied upon here would always find instances of poor performance that had not yet been punished by the market. Finally, there are almost certainly measurement errors and other problems in the WCC price and quality data that could inject random noise and prevent precise measurement of the actual level of price-quality correlation in the market.⁵

At the other extreme, it is unlikely that any of the surveyed markets are completely "imperfect" in the sense that consumers are totally incapable of gathering any price or quality information whatever. Thus, about all we can predict about the absolute correlation between price and quality that we might observe from the WCC data for any given industry is that it should be greater than zero and less than one.⁶

Fortunately, the WCC data and existing theories from the literature on the economics of information allow us to test other interesting hypotheses concerning the expected comparative strength of price-quality correlations across service industries, and to test the accuracy of certain non-price "signals" that consumers might use to assess the quality of individual providers within industries. The availability of time series data for most of the sampled industries also permits an analysis of the stability of a firm's price and quality ratings over time. Specifically, it is possible to test whether a firm that receives a high rating for quality in an initial period is likely to "milk" any reputation gains and lower quality or raise its price in subsequent periods.

A. Price-Quality Correlations Across Industries

The first area for exploration is the pattern of price-quality correlations that we observe from industry to industry. Theory would predict that these correlations will be highest in those service markets where consumers can compare the prices and quality of competitors at low cost. It is relatively easy to specify the characteristics of a service that would determine how costly it would be for consumers to shop for price. Search will be easiest when the desired service is standardized, or comprised of standardized subcomponents, and the full scope of the needed service is known to the consumer before purchase. These characteristics would allow the consumer to gather price information efficiently by word-of-mouth, telephone, advertisements, or

⁵ We will ignore for purposes of this discussion the possibly perverse bias introduced by any interaction between prices and consumer satisfaction that could reveal a seemingly negative price-quality correlation.

⁶ Indeed, we would not necessarily expect a simple correlation of one in a perfectly functioning market. Although quality should increase monotonically with price under such circumstances, the relationship would not be linear unless the long run marginal cost of producing additional quality happened to be perfectly linear.

on-site inspection. The clearest examples of such services would be drycleaning, carpet cleaning, restaurants, and supermarkets.

Almost by definition, consumers cannot directly determine a service provider's quality until after the service is rendered. Still, there may be wide variance among industries in the ease with which quality can be evaluated post-purchase. In certain cases, quality can be evaluated fully by simple inspection or immediate experience. Again, restaurants, carpet cleaning, and drycleaning score well in this regard. In such instances, the quality reputation of firms would spread fairly quickly and allow new consumers to locate the desired combination of price and quality at minimal cost.

At the other end of the spectrum, comparative price and quality comparisons will be most difficult where the service is complex and the consumer may not know pre-purchase precisely which service components will be needed, or post-purchase whether the service had been fully or honestly rendered. Examples would include certain repair services for automobiles, appliances, and electronic entertainment items. The full cost of the repair would not be known until the provider had diagnosed the problem, at which point the consumer may already have incurred substantial out-of-pocket expenses and might be reluctant to repeat the process with another establishment.⁷ Although consumers could determine post-purchase whether a repair had fixed the problem, they might not know whether all of the billed services were necessary.

Though not strictly a determinant of search cost, another factor bearing on price-quality correlation should be the relative frequency with which consumers purchase a given service. Clearly, consumers will have more opportunity to sample and compare competitors' offerings the more often they shop. For local services, this means (all else equal) that the market should punish poorly performing drycleaners more quickly than roofers who offer poor quality for the money. Although it was hoped that relative frequency of purchase could be quantified using existing marketing data or information supplied by the principal trade associations for each sampled industry, such data apparently are not available. Thus, very broad and subjective assessments of this and other consumer search costs factors are used to test whether price-quality correlations are in fact higher in low-search-cost industries.

B. Price-Quality Relationships Within Industries

The preceding analysis has focused on informational and purchase pattern considerations that might explain differing degrees of price-quality correlation across local service industries. It

⁷ Comparative shopping will also be difficult or impractical when service is required on an emergency basis. There are many examples of service industries that may respond to emergency situations (e.g. Plumbers, Roofers, Air Conditioning and Heating firms, Auto Repair, and, for entertainment-challenged families, TV Repair), but this would not be the exclusive or even primary form of business for such firms. It is therefore difficult to sort industries on this basis.

would also be of interest to test theories that might predict which types of firms *within* a given industry are likely to offer higher levels of quality, or perhaps higher levels of value. That is, in those markets where consumers do not believe they can rely upon price as an accurate indicator of firm quality, are there other firm characteristics or practices that consumers can use as a time and money-saving substitute for extensive search in determining the level of quality and price that a firm is likely to provide?

1. Signaling Mechanisms

This issue has been treated extensively in the "signaling" literature, which originated in 1974 with Spence [15] and Nelson [11]. All else equal, it is hypothesized that higher quality providers will have a greater incentive to signal their advantage through advertising and through other bonding devices such as nontransferable physical capital (e.g., specialized building designs and accoutrements), the value of which is lost if consumers are disillusioned with a provider's quality post-purchase [8, 9]. In such cases, the advertising or other signal need not have any informational content *per se*. It is the mere existence and size of the advertising or other nontransferable physical investment that assures consumers that the firm is trustworthy and prepared to please customers over the long term.⁸

In some respects, service industries do not appear to fit the signaling model very well. First, with the exception of national chains and dealer franchises present primarily in the auto repair industry (discussed separately below), few of the firms in the industries rated by WCC engage in broadcast or other major media advertising. Many of the firms serve smaller regional markets and do not operate at a scale that would justify significant advertising outlays, although such firms may place ads in smaller regional newspapers and church bulletins.

In addition, firms in many of the industries under examination provide their services in the consumer's home. Thus, bonding signals in the form of fancy, imposing office sites would not prove very effective. Many of the remaining service industries perform repair functions, where the physical plant is likely to be a basic facility readily transferrable to other uses. There is, however, one form of advertising in this sector that might perform a signaling function. This is Yellow Pages advertising.

a.. Yellow Pages Advertising

For many consumers, the Yellow Pages will provide the first and sometimes only step in the search process. Firms that wish to stand out in the crowd of listings must pay a substantial

⁸ The signaling literature is exceedingly complex, and the various signaling models often incorporate critical assumptions specifying precise consumer knowledge of firms' marginal and fixed costs. If these assumptions are relaxed, a positive relationship between quality and advertising or other signals need not be reached in equilibrium. In the words of a recent survey article, "Anything can happen." [5], p. 34.

premium for full and half-page ads. Specifically, a full-page ad in the Suburban Maryland Yellow Pages costs about \$42,000 a year.⁹ Reaching the entire Washington D.C. area with full-page ads would cost over \$100,000 annually.

Although Yellow Pages advertising does incorporate some of the features of a signaling mechanism, it is not a particularly strong test of the theory. Specifically, in a community with high population turnover such as the greater Washington, D.C. region, it is not obvious that the cost of conspicuous Yellow Pages ads would be so high as to require extensive repeat business to justify the investment. The payoff from attracting more first-time, and only-time, purchasers who are new to the region may well be sufficient to warrant such advertising. Indeed, short-term gains may be high enough even in the absence of high population turnover.

In addition, heavy outlays for a large Yellow Pages ad would not be a feasible or even desirable strategy for many firms in the local service sector economy. The service industries at issue are a mixture of small regional providers who serve only a single jurisdiction or even neighborhood, and area-wide firms with perhaps many outlets located in Virginia, Maryland, and the District. Although both types of firms clearly compete with one another, most of the smaller providers may purposefully limit their scale of operation. Such small, localized firms may have no interest in running a very large Yellow Pages ad. Rather, their marketing strategy would be to rely on repeat purchases from current clients and word-of-mouth advertising in their local market area, and to confine their Yellow Pages presence to a single line listing. (In the statistical analysis reported below, the relationship between Yellow Page advertising and quality is explored using both the full sample of firms and a subset of firms that excludes single line listings.)

b. Franchise and National Chain Status

As mentioned previously, a small subset of the studied industries contains firms that are franchisees of a national manufacturer or are members of a national chain. Franchised dealerships are most common in the auto repair sector. There are also several national auto repair chains represented in the Washington, D.C. area (*e.g.* Goodyear, Merchant Tire, Sears, and Jiffy Lube.) National chains or dealerships are present to a lesser extent in the auto body repair, pest control, and major appliance repair industries.

National chains, such as Sears and Jiffy Lube, advertise extensively and enjoy considerable brand name recognition. The non-informational component of these advertising investments can be construed as a pure signal to consumers that the local chain representative

⁹ This information was provided to the author by Marc Rysman from a data set containing pricing information for almost all Yellow Pages directories in the United States in 1997. See M. Busse and M. Rysman, "Competition and Price Discrimination in Yellow Pages Advertising" Working Paper Series ES, Number 13, Yale School of Management, June 2001.

will satisfy its customers, either in terms of the absolute quality of the service or in terms of quality adjusted for price.

Franchised car dealerships (*e.g.* Toyota, BMW, etc.) may benefit from any goodwill associated with the automobiles they represent, and consumers may assume that dealers have special expertise in repairing their brands of cars. (It should be noted, however, that the potential quality clue offered by franchise status does not constitute a pure signal in the economic sense, since its utility rests on direct information concerning the known quality of the automobile brand and logical inferences concerning the likely expertise a dealership would gain from specializing in the repair of a limited number of automobile brands.)

The hypotheses that chain or dealership affiliation is a reliable indicator of quality or value can be readily tested with the WCC data, since such affiliations are obvious from the title of the firm. Separate zero-one dummy variables were constructed for dealer and chain status, and included as a right-hand term in regressions using WCC quality ratings as a dependent variable.

2. Firm Size

Many consumers may also make quality inferences based on a service provider's size. Size is not a pure signal in the sense described above, since size is a highly complex market outcome rather than a simple short-term investment decision. Further, the relationship between size and quality (as well as price) is an interesting and important economic issue irrespective of whether consumers know a firm's scale of operations or make use of any information they might have. The fundamental question, of course, is whether large size is in fact a market reward for superior performance.

Firms might prosper simply by providing above-average quality at a higher-than-average price, assuming there was a substantial market for such a combination of price and quality. Alternatively, firms may attempt to grow by exploiting any scale economies in such activities as purchasing, inventorying, certain administrative functions, and scheduling. A priori, it is not clear what absolute level of quality would be optimal for such firms. Presumably, however, these providers would attempt to obtain increased market share by passing at least some of their cost savings through to consumers, and we would therefore expect the price charged for any given level of service quality to be below the industry average for that quality level.¹⁰ The WCC data allow us to test whether larger firms offer higher quality, and/or whether such firms offer higher value as measured by consumer quality ratings adjusted for price.

¹⁰ This is the outcome that <u>Bond *et al.*</u> [3] observed in their analysis of optometrists operating in states that did and did not allow price advertising and large chain operations. In nonrestrictive states, advertisers provided slightly less thorough eye exams than did smaller nonadvertisers, but charged substantially less for the bundled combination of eye glasses and eye exams.

3. Third Party Certification and Endorsement

Official authorizations, accreditation, or certification from various private or public sources may provide consumers with very straightforward indicators of firm quality. For the local services in this data set, such third party certifications are most prevalent in the auto repair industry. Examples of private endorsements include approval by the American Automobile Association and certification of mechanics by the National Institute of Automotive Service Excellence. Consumers might also infer a quality advantage for repair shops that are approved by local or state governments for safety and/or emissions inspection and testing. The WCC data contain shop-specific information on such certifications and authorizations, and thus provide one means of testing their usefulness to consumers as quality clues.

C. Price and Quality Performance Over Time

As previously discussed, WCC has evaluated most of the 19 industries included in our sample on several occasions since 1976. The time lapse between evaluations ranges from a minimum of three years to over a decade for the infrequently rated industries. These various snapshots in time allow us to evaluate the stability of an individual firm's price and quality ratings and to measure any trends in an industry's overall performance from its initial assessment to the present. Finally, the availability of multiple data sets for most of the sample industries provides a consistency check for the statistical analysis of price-quality correlations and certain other relationships.

III. DETAILED DESCRIPTION OF DATA SET AND VARIABLES

The data set is comprised of 19 local service industries, which represents every industry with WCC price information and consumer quality ratings. As reported in Table I above, in most instances separate data sets were constructed for each evaluation an industry has received since the magazine began publishing in 1976. There are a total of 60 such data sets.

A typical data set for a given industry would include the following variables taken from the WCC ratings: (1) a price index for each firm as constructed by WCC; (2) the number of complaints on file for a given firm at local consumer protection offices, in certain cases adjusted for the size of firm; (3) the quality rating given by the survey respondents; (4) the number of respondents rating a firm; and (5) two zero-one dummy variables to indicate whether WCC had check rated the firm for quality and/or price.

The respondent quality rating variable requires further elaboration. The survey instrument used by WCC asks consumers to rate a firm on the basis of "overall performance," "doing work properly the first time," and a variety of other performance dimensions, such as

"letting you know cost early," "starting and completing work promptly," and in some cases, even "neatness." Unless otherwise noted, all of the results reported below are based on the "overall performance" category. Although the "doing work properly" category is in some sense the most focused measure of a firm's pure competence, excelling in the other dimensions of performance included in the "overall performance" category should impose costs on a firm, and these attributes of performance properly should be viewed as components of quality. (In all cases, however, sensitivity tests were conducted to determine whether an alternative measure of quality would alter the results substantially.)

Within each category of firm performance, consumers may rate a firm as "superior," "satisfactory," or "unsatisfactory." For a majority of industries, WCC reports for at least the "overall performance" category both the percentage of respondents rating the firm superior, and the percentage rating the firm either satisfactory or superior. For six of the industries, however, only the broader "satisfactory or superior" percentage is provided. This omission reduces the variance in the quality ratings and could reduce any explanatory power the independent variables might have.

In addition to the WCC variables, supplementary information was gathered for the most recent rating period for 14 of the industries to test for any signaling function provided by Yellow Page advertising. The resulting variable is simply the size of a given firm's advertisement measured in square inches, with a one line regular print listing coded as .25 square inches, and a bold listing coded as .5 square inches. Thus, this variable has a potential range of .25 to 80 for a full ten-by-eight-inch page.

The WCC data set was also supplemented in certain cases with information on firm size as measured by number of employees. WCC was able to supply this information for a total of seven industries. Finally, dummy variables were created where appropriate to represent whether a firm was a dealership or a franchisee of a national chain. Certain other dummy variables were used in the Auto Repair data set, which is discussed separately below.

IV. STATISTICAL METHODOLOGY

For most of the regression analysis reported below, the left-hand variable is the WCC overall performance rating, which is the percentage of survey respondents rating a firm superior (or, in some cases, either satisfactory or superior). There are a number of statistical issues involved in use of this measure as a dependent variable. First, the measure can only vary from zero (percent) at a minimum to 100 at a maximum. Further, a substantial number of summary ratings cluster at or near the upper end of the scale. Under such circumstances, ordinary least squares is not a suitable estimation technique, since the error terms will not be distributed normally.

In addition, the error terms for observations within any given firm will be correlated, since there will be variables not included in our analysis that will have an impact on the overall level of satisfaction ratings for each firm. One such factor could be the politeness and general demeanor of the staff. Firms faring well in this dimension will receive higher WCC quality ratings, which means that a regression that did not capture this determinant of satisfaction would systematically underestimate the quality scores reported by respondents for such firms.

Finally, the number of respondents rating a particular firm varies widely within industries, from a minimum of ten to a maximum of 500 or more. Since these ratings are in effect sample estimates of the true ratings that all customers of that firm would give, the precision of the estimates increases with the number of ratings. It is appropriate, therefore, to weight firm quality ratings by the number of individual ratings on which they are based.

All of these econometric issues were addressed by adopting a logit estimation technique that is structured for use with "grouped" data. In essence, this procedure unfolds the summary quality rating for a given firm into a series of binary categorical variables, with a zero-one value generated for each respondent rating used in calculating the overall quality score. Consider, for example, a firm that receives an overall performance rating of 90, which for most industries would mean that 90 percent of the consumers reporting on that firm rated it superior. If there were a total of 10 respondents, our logit technique would generate a series of dependent variables comprised of nine "one" values and one "zero" value. Each of these dummies would be associated with the corresponding firm-specific independent variables used in the logit regression. Thus, firms with the highest number of individual ratings automatically would be weighted heaviest in the regression estimation procedure. This procedure also adjusts standard errors to correct for the expected correlation of error terms within firms.

V. SUMMARY OF PRINCIPAL FINDINGS

A. Overview of Quality Ratings

Table II below presents the average WCC overall quality ratings for the various industries in the sample for the most recent survey period. The first rating is the average percentage of respondents that gauged the overall performance of firms in the industry as superior. The second rating combines the superior percentage with the percentage of respondents that rated a firm as "adequate." WCC did not disclose ratings for the superior category in six industries, and in four other industries did not provide the combined score for superior and adequate. In the case of restaurants, respondents were asked to rate the quality of the food on a scale of zero to100. The resulting quality measure therefore is not comparable to the average percentage score ratings for the other industries. Industries are ranked in ascending order by their average rating in the "% Superior or Adequate" category (where available).

Year Rat	ed Industry	% Superior	% Superior or Adequate
98	Supermarkets	36.0	
96	Carpet Installers		80.2
98	Computer Repair		80.2
89	Television Repair		87.7
01	Auto Repair		89.1
96	Appliance Repair		89.8
99	Pest Control	59.6	90.5
98	Auto Body Repair	69.3	90.5
96	Watch Repair	74.5	91.1
98	Plumbers		91.8
97	AC & Heating	70.7	92.9
96	Drycleaners	61.4	94.0
98	Local Movers	69.4	
98	Carpet Cleaning	73.7	
96	Electricians	78.2	94.4
99	Locksmiths	79.5	95.8
95	Shoe Repair	81.0	
99	Tree Experts	79.9	96.5
98	Restaurants		75.4*

Table IIWCC Quality Ratings for Sample Industries

*Restaurants are rated on a scale of 0-100

Table II reveals a few interesting patterns and conclusions. First, using the more sensitive "% Superior" rating, Washington area customers are clearly not extremely enthusiastic about the overall quality of local supermarkets. It is not clear whether this reflects an underlying problem, or merely the extreme familiarity consumers have with the food shopping experience. And, for whatever reason, local consumers do seem enthusiastic about their choice of shoe repair firm. Between these two boundaries, industries specializing in mechanical repair do not fare as well as the more traditional "guild" industries, such as electricians, locksmiths, and, as mentioned, shoe repair.

On an overall basis, the various ratings suggest that the local service sector is functioning fairly well from an absolute standpoint. An average of 90 percent of consumers rated their service provider as at least adequate, and an average of 69 percent graded the performance as

superior. At a minimum, these percentages do not suggest a major market meltdown of the magnitude associated with a "lemons" model outcome.

B. Consistency of WCC Quality Ratings and Complaint Data

As discussed, the WCC ratings list the number of complaints on file at local consumer protection offices for each rated firm. For certain industries, this number is normalized for firm size, as measured by number of employees. The complaint variable was included in the initial regression runs to function as a possible consistency check on the WCC quality ratings. Since a firm's complaint history is really an indirect measure of quality, it cannot be construed properly as a true independent variable in any regression that uses the WCC quality ratings as a dependent variable. It does have the potential, however, to shed some light on the reliability of the respondent ratings.

In particular, the complaint variable can help determine whether the WCC quality ratings are hopelessly biased by the possibly perverse impact that a firm's prices might have on consumer attitudes toward the firm. As described earlier, this interaction between price and consumer satisfaction, combined with possible selection biases, might yield a strong negative correlation between price and quality, even when true quality was in fact directly related to price.

Consumers presumably take the trouble to write complaint letters over what is perceived as truly egregious firm behavior. Although these complaints may involve what a consumer feels is an exorbitant price, complaints should not be generated by the more subtle impact that, say, a somewhat above-average service repair charge price might have on the disposition of a WCC respondent rating the overall performance of a firm. Thus, the complaint rate variable should help flag firms that are genuinely poor performers, and this variable should be negatively associated with the WCC ratings if these ratings are at all reliable.

The results indicate that there is in fact a persistent and strong negative correlation between firm complaint histories and the quality ratings. For the most recent rating periods, WCC provided some form of complaint data for 14 of the 19 industries. In ten cases, the data were adjusted by firm employee size. In seven of these industries, the complaint rate was negatively and significantly correlated with quality. In the three remaining industries, the coefficient sign was negative but not significant. For five industries, WCC provided only the absolute number of complaints a firm had on file. The complaint coefficients for three of these industries were negative and significant, and insignificant in the remaining two. In the entire data base, the complaint rate variable was negative and significant in approximately 90 percent of the industries for which size-adjusted data were provided. These results, though far from conclusive, do suggest that the WCC ratings are at least flagging the very worst performers.

C. Simple Price-Quality Correlations

In this section, we discuss the threshold question of whether price by itself provides a reliable indicator of quality as measured by WCC respondent performance ratings. The regressions reported do not control for any firm characteristic, such as size or status as a franchisee or dealership, or any signaling mechanism, such as Yellow Pages advertising. The analysis simply attempts to determine the extent to which consumers can rely upon price alone as an indicator of firm quality.

The results of the simple quality-price regressions for the most recent ratings period are reported in Table III. The coefficients were obtained by grouped logit analysis using the WCC quality rating as the dependent variable. P values for the price coefficients appear in parentheses.

Industry	Coefficient Sign for Price (P-value)
Carpet Cleaning	Positive (.009)
Carpet Installers	Positive (.090)
Dry Cleaning	Positive (.102)
Local Movers	Positive (.216)
Pest Control	Positive (.000)
Restaurants	Positive (.000)
Supermarkets	Positive (.000)
Tree Experts	Positive (.403)
Auto Body Shops	Negative (.548)
Air Conditioning-Htg.	Negative (.034)
Appliance Repair	Negative (.000)
Auto Mechanical Repair	Negative (.000)
Computer Repair	Negative (.078)
Electricians	Negative (.001)
Locksmiths	Negative (.003)
Plumbers	Negative (.007)
Shoe Repair	Negative (.131)
Television Repair	Negative (.307)
Watch Repair	Negative (.732)

TABLE III Quality-Price Correlations in Most Recent Rating Period

As hypothesized, those industries with particularly low search costs and relatively frequent purchase generally display significant positive correlations between price and quality. These are Carpet Cleaning, Restaurants, Supermarkets, and, just missing significance at the .10 level, Dry Cleaning (P=.102). In addition, price and quality are also positively correlated for Carpet Installers and Pest Control. The most striking feature of Table III, however, is the relentless procession of negative correlations in the eleven repair industries, seven of which are significant.

The picture does not change substantially when results for all of the ratings periods are considered. Table IV provides a detailed listing of the observed price-quality correlations for each of the 60 data sets in the sample. Table IV reveals significant positive correlations in ten data sets, representing seven industries. Of these industries, however, only Carpet Cleaning, Drycleaners, and Restaurants are consistently positive over time. Supermarkets reports a highly significant positive price-quality correlation in the two later ratings periods, but is insignificant in 1979. Significant negative price-quality correlations can be observed in 22 data sets, accounting for ten industries, and 28 sets display no significant price-quality correlation.

The widespread occurrence of negative price-quality correlations is difficult to rationalize using any rigorous theory of market performance. Even in the presence of extremely high search costs, there are no *a priori* grounds for expecting price to serve as a *perverse* indicator of quality. Rather, we would simply expect a great deal of noise with no systematic relationship between price and quality. If we are to believe the results reported above, firms apparently prosper by choosing a strategy of high prices and poor performance. It is implausible that even serious market imperfections would perpetually reward such a strategy.¹¹ Thus, we certainly cannot dismiss the hypothesis that the WCC data are biased due to the previously discussed interaction between price and consumer satisfaction with firm performance.

One alternative hypotheses for certain industries in the sample is that the simple quality and price regressions fail to control for relevant firm cost variables, particularly higher rental

¹¹ Such an outcome might occur under the extreme assumptions that consumers never purchase a service more than once and that absolutely no quality information is available prepurchase. In that event, some firms might successfully pursue a strategy of high price and low quality, since consumers seeking high quality might specifically target firms quoting the highest price for a service, and by assumption such firms would never be punished for failing to deliver the expected quality. Once these assumptions are relaxed to allow for at least some repeat purchases and limited availability of quality information, it is once again difficult to understand how the high price-low quality strategy could persist so consistently in so many industries over such a long period of time.

Industry	Year	Sign	P Value	Observations
AC & Heating Contractors	97	negative	.034	119
C C	92	negative	.273	106
	87	negative	.428	78
	82	negative	.518	61
	77	positive	.972	44
Appliance Repair	96	negative	.000	46
	91	negative	.002	43
	85	negative	.012	52
	78	negative	.228	60
Auto Body Repair	98	negative	.548	120
	95	negative	.422	106
	90	negative	.393	90
	85	negative	.191	105
	82	negative	.041	71
Auto Mech. Repair	01	negative	.000	482
-	97	negative	.000	444
	94	negative	.000	431
	91	negative	.000	380
	88	negative	.000	354
	85	negative	.000	310
	81	negative	.000	289
	76	negative	.004	148
Carpet Cleaning	98	positive	.009	33
	94	positive	.425	31
	87	positive	.727	30
Carpet Sales, Installation	96	positive	.090	24
	89	negative	.723	39
Computer Repair	98	negative	.078	27
	94	positive	.248	31
	89	negative	.929	14

Table IV
Quality-Price Correlations for All Industries and Ratings Periods

Industry	Year	Sign	P Value	Observations
Drycleaners	96	positive	.102	211
	90	positive	.038	232
Electricians	96	negative	.001	36
	90	negative	.036	21
Locksmiths	99	negative	.003	25
	94	negative	.499	27
	86	negative	.587	33
Movers	98	positive	.216	28
	92	negative	.720	29
	81	positive	.051	25
Pest Control	97	positive	.000	51
	93	positive	.477	37
	86	negative	.292	46
	77	negative	.030	25
Plumbers	95	negative	.007	134
	89	negative	.045	81
	83	negative	.019	97
	77	positive	.165	58
Restaurants	98	positive	.000	672
	95	positive	.000	718
Shoe Repair	95	negative	.131	95
	88	positive	.213	96
Supermarkets	01	positive	.000	8
	92	positive	.079	7
	79	positive	.997	16
Television Repair	89	negative	.307	36
Tree Experts	99	positive	.403	29
Watch Repair	` 96	negative	.691	41
	91	negative	.181	30
	80	negative	.002	52

Table IV (Continued)

costs in more convenient and/or affluent locations. Consumers presumably would willingly pay more for a tune up in a shop or dealership located close to work or home than one in an industrial park in an outer suburb. Thus, the simple correlations potentially suffer from a missing demand variable in the form of convenience, and a missing supply variable in the form of rental costs. This defect would affect a fairly small number of the sample industries, since so many of the services studied here are performed in the home and do not involve consumer visits to provider sites.

In an attempt to remedy this problem, census household income data by census tract were collected and used as an independent variable in a separate price equation for the relevant industries.¹² Although direct rental cost information by, say, zip code would have been preferable to the income variable, these data were not available. It was hoped that the income data would be sufficiently correlated with rentals to provide a satisfactory substitute. For whatever reason, however, none of the price equations produced significant results, and no subsequent attempt was made to estimate a reduced form equation that would control for convenience and rental costs.

D. Tests of Signaling Theories and the Performance of Larger Firms

Our data set allows tests of several signaling mechanisms, and also a test for whether firm size is associated with quality. The variables and hypothesized outcome for size of Yellow Pages ad, dealership status, and franchisee status have been discussed earlier. In addition, the Auto Repair and Pest Control data sets permit certain other signaling tests that will be discussed separately. Results were obtained from grouped logit regressions that used only the independent variable under examination as a predictor of quality. This specification is appropriate to test the hypothesis that consumers can rely upon the variable of interest in isolation as a signal of quality. See Appendix A for a presentation of the full regression results underlying the discussion in this section.

1. Yellow Pages Advertising

Data on the size of Yellow Pages ads were collected for the most recent rating period for 14 of the 19 industries. This variable was used in several specifications to test for robustness and any extreme value problems that could be introduced by the very wide range in values (from .25 to 80), and the possibility that one or two firms with full page ads might by chance carry particularly high or low quality ratings. Specifically, additional regressions were run with the Yellow Pages variable in log form, and subsets of firms were tested to determine whether any relationship between quality and advertising was limited to advertisements above or below a certain size. In particular, runs were always made excluding firms with one-line listings. As

¹² These industries were Auto Body Repair, Auto Mechanical Repair, Drycleaners, Shoe Repair, and Watch Repair. Due to resource constraints and the strong positive price-quality correlation already evident, data were not collected for Restaurants.

discussed earlier, such firms might deliberately have chosen to limit their scale of operation and would never have considered signaling quality using a larger Yellow Pages ad.

The results provide little support for a Yellow Pages signaling function. A significant positive coefficient was found for only four of the 14 industries. In seven industries, the Yellow Pages variable was negatively and significantly associated with quality. There was no significant relationship in the remaining three industries. Simple regressions were also run to investigate whether the size of an ad might signal higher value as measured by the ratio of the quality rating to the price index. None of these regressions supported this hypothesis. Indeed, a significant negative correlation between Yellow Pages ad size and value was found in five industries.

2. Dealership and Franchisee Status

In two industries, Auto Mechanical Repair and Auto Body Repair, numerous firms in the sample were franchised dealerships for the major auto companies. In four industries (the two auto repair industries, Local Movers, and Pest Control), many of the firms were franchisees for national chains, such as Maaco, Sears, United Van Lines, and Orkin Exterminating. In these industries, dummy variables were used to test whether consumers rated such firms higher than independent firms. The results were uniformly negative. The coefficients on these variables were always negative, and usually highly significant. (The results for Auto Mechanical Repair will be presented in more detail below.) Although franchised dealerships consistently charged more than non-dealers, the prices of national chains did not differ significantly from those of independent providers.

3. Firm Size

The firm size variable is of interest for two reasons. First, it clearly would be interesting to see whether larger firms might have attained their size advantage because they provided superior quality. Second, it is of interest to determine whether larger firms are realizing scale economies that might be passed on to consumers in lower prices for a given level of quality. For this data set, the answer to both questions is a resounding no. WCC provided employee size information for seven industries, and in all seven industries there is a very significant negative correlations between quality and size. Nor did larger firms provide greater value in terms of quality adjusted for price. The coefficient on the value variable was negative in all cases, and significantly so in five industries.

4. Warranties

One of the industries–Pest Control–provides a unique opportunity to test whether the length of warranty protection that firms provide signals higher reliability, or whether the warranty functions purely as indemnification in the event of actual product failure. The WCC ratings for pest control firms contain a detailed breakdown of firm charges and terms for a pest eradication treatment for a typical house. WCC lists the initial charge for the treatment, the time period

during which the firm will provide a free follow-up treatment, and the charge for treatment after the free follow-up warranty has expired. The warranty varies widely in the sample, from zero days to a full year. Further, as is the case for almost all industries, WCC provides customer ratings for a firm's ability to do the work properly on a first try, as well as an overall performance assessment.

If firms offering longer follow-up warranties were more competent and diligent, so that consumers could rely on the warranty period as an indicator of quality, one would expect that such firms would receive higher ratings for doing the initial treatment properly. This is not the relationship we observe in the Pest Control data set. The rating for doing work properly tends to be negatively and significantly associated with the length of the follow-up treatment warranty.

5. Certification

Finally, the Auto Repair data set contains information that can be used to test whether consumers can rely on various forms of certification and authorization as indicators of quality. WCC indicates for each auto repair establishment whether it is approved by the American Automobile Association, whether the firm employs at least one ASIE technician, and whether it is licensed to perform state safety inspections. The first two variables provide straightforward quality certification tests. The hypothesized coefficient for state safety inspection status is less clear, since many repair establishments may question the profitability of such inspections, and may decline participation for reasons unrelated to provider competence.

Certification by AAA is positively but not significantly correlated with the WCC ratings in the seven auto repair data sets. The ASIE variable, on the other hand, is frequently negative and significant, and never positive and significant. State Safety Inspection facilities consistently fare worse in the quality rankings.

As presented in Appendix A on pp. 43-44, the most dramatic result from the simple regression runs is the strength of the negative correlation between the dealership and chain variables and overall performance. In simple numeric terms, without controlling for any other variables, dealerships on average received an overall satisfaction score of 80.6 in the 1997 ratings, 12 points below the rating of 92.6 for independent shops. Chains, with an average rating of 76.9, scored even lower than dealerships.

Firm size also displays a highly significant negative correlation with quality. Since dealerships and chains tend to be larger than independent shops, it might be concluded that the size variable is merely serving as a proxy for dealership and chain status. As shown in Table V below, however, multivariate regression analysis demonstrates that the significant negative coefficient on firm size persists when dealership and chain status are also included as independent variables.

Further, this negative association between size and quality or between chain status and quality cannot be attributed to any interaction between price and the consumer satisfaction

ratings. Price is not correlated with either firm size or chain status in this data set. (Dealerships, however, charged significantly higher prices than independents or chains.) Thus, the negative coefficients for the size and chain variables apparently should be accepted at face value.

Dependent Variable = % rating firm superior or adequate								
Logit estimate	S	Numbe	r of firms =	441.00				
LR chi2(7) = Prob > chi2 = Log likelihood = -6423.8417 Pseudo R2 =								
supq	Coef.	Std. Err.	z	P> z	[95% Conf	. Interval]		
chain97 dealer97 Price97 Size7 ASIE97 AAA97 Safety97 Constant	-1.31156 9997785 .0046225 0147877 8662402 .189219 3748802 3.403598	.0806641 .0015918 .0034198 .2126259	$ \begin{array}{r} -13.47 \\ -12.39 \\ 2.90 \\ -4.32 \\ -4.07 \\ 3.43 \\ -6.12 \\ 13.83 \end{array} $	$\begin{array}{c} 0.000\\ 0.000\\ 0.004\\ 0.000\\ 0.000\\ 0.000\\ 0.001\\ 0.000\\ 0.000\\ 0.000\\ \end{array}$	-1.502361 -1.157877 .0015025 0214904 -1.282979 .0812489 4948866 2.921095	.0077424		

TABLE VFull Regression Results For Auto Repair 1997

It should also be noted that the price index variable is positively and significantly correlated with quality in Table V, whereas the coefficient was negative and highly significant when used as a single predictor. The primary explanation for the shift in signs is the explicit accounting for dealer performance in this regression. As discussed, dealers tend to be more expensive and do considerably worse in the ratings than independents. Thus, price acts partially as a proxy for dealership status when used alone as a predictor of quality.

The positive correlation in the full regression cannot, however, be interpreted as an indicator of market efficiency. The reversal in sign merely indicates that we have identified and controlled for one possible source of poor industry performance. On the other hand, the seemingly poor performance of dealers may itself be an artifact of possibly strong selection biases associated with dealership status. For example, if consumers seek out dealers for particularly complicated repair tasks, the probability of service problems inevitably would be higher and reflected in the lower satisfaction ratings.

E. Stability of Price and Quality Ratings Over Time

In this section, we first explore any overall trends in industry quality performance since WCC began its ratings in 1976. (The price data are virtually always normalized, with an average price within any industry assigned a value of 100, and therefore are not suitable for trend analysis.) We then turn to the individual firm level and analyze the stability of firm performance through ratings periods. In particular, we test whether firms that are singled out for a check rating for price or quality tend to maintain a check rating in subsequent periods.

1. Overall Industry Trends

It is of some general interest to determine whether overall performance in the sample service industries has improved over the years, due perhaps to technological improvements or even to the impact of the WCC ratings themselves. There is, unfortunately, a potentially incestuous influence in the WCC data that will tend to improve average satisfaction scores irrespective of any actual increase in quality among all the firms in the industry. Once an industry has been evaluated in an issue of WCC, readers of the magazine can be expected to patronize firms that do well in quality and/or price. In the subsequent rating period, firms that scored poorly initially will tend to receive fewer ratings as readers shift their allegiances to more favored providers, and some will drop out of the sample entirely due to inadequate response rates. Thus, there should be a built-in upward trend in the average quality ratings.

With this proviso, we present in Table VI the results of a simple trend analysis obtained by regressing the average overall satisfaction scores for a given industry on a time counter. In some cases, the number of ratings periods is very low and the resulting "trend" is not particularly interesting. For what they are worth, the results show that, for the 16 industries with more than one WCC rating period, overall satisfaction increases significantly in nine, decreases significantly in four, and displays no trend in three.

TABLE VI Trends in Service Industry Consumer Satisfaction Over Time

Significant Upward Trend (P<.10): Auto Body Repair, AC & Heating Contractors, Major Appliance Repair, Auto Mechanical Repair, Drycleaning (2 periods), Electricians (2 periods), Locksmiths, Shoe Repair (2 periods), Watch Repair.

Significant Downward Trend: Local Movers, Pest Control, Plumbers, Restaurants.

No Significant Trend: Carpet Cleaning, Carpet Installers, Computer Repair.

2. Individual Firm Performance Over Time

Of considerably greater interest than overall industry trends is the issue of whether firms that are singled out for special recognition by WCC for either price or quality tend to repeat their superior performance in the subsequent ratings period. It is possible that some firms might attempt to exploit a quality check rating by, say, lowering costs and quality while maintaining or even increasing price. Similarly, a firm spotlighted for particularly low prices might seek short run gains by increasing prices in the hope that a substantial number of consumers acting on the rating might not realize that the firm had lost its price advantage.

Ideally, we would like to track firm performance shortly after the ratings appeared in order to see the initial impact on what might prove to be short term firm behavior. Because of the long time lag between ratings periods, however, the WCC ratings only allow a check after several years have elapsed. We therefore are limited to assessing any possible direct or indirect long term impact of the ratings on firm performance.

This issue was explored using probit analysis. Specifically, we tested whether firms that were check-rated for price or quality in a given rating period have a greater-than-average probability of being check-rated again in the subsequent period. For any industry with N rating periods, N-1 probits were run so that all subsequent ratings were considered in sequence.

In order to give more intuitive meaning to the probit coefficients, the results are reported in terms of marginal probabilities, which show the *difference* in the probability that a check rated firm vs. a non-check rated firm will receive a check rating in the subsequent period. That is, a marginal probability of .42 indicates that a firm check rated for, say, price in the first period will be 42 percentage points more likely than a non-check rated firm to receive a price check rating in the next period.

The probit analysis provides a more rigorous and meaningful test of any tendency for firms to "milk" reputation effects than would simple comparisons of a firm's price index score or its quality rating between periods. For any given rating period, some firms may tend to score well on the quality survey or price index purely for stochastic reasons. By luck of the draw, an otherwise average firm might find that it had alienated or pleased a particularly large number of consumers who happened to have received the WCC survey form. Similarly, WCC might by happenstance have chosen a mix of service tasks for its price index that particularly advantaged or disadvantaged a firm because of atypical circumstances. A plumbing contractor, for example, might have been short on skilled personnel temporarily on the day WCC requested an estimate for some nontrivial task, and may have quoted a noncompetitive price.

When such firms were rated again in the next period, such stochastic events would tend to even out, and the firms's ratings would tend to "regress to the mean." Thus, to some extent, we would always expect to find on average that firms check rated for, say, quality in 1992 would score somewhat more poorly in 1998. Similarly, an average firm in 1992 might benefit from

chance events and earn a check rating in 1998. We could not conclude from this pattern, however, that check rated firms tend to milk their reputations and lower quality.

The probit analysis controls for such random fluctuations and shows more clearly whether there are real differences in the propensity for check rated and non-check rated firms to score well in the subsequent period. Still, the results do not lend themselves to an unambiguous test of the milking hypothesis. In the example above, a significant marginal probability of .42 would indicate that there is no pervasive tendency for firms to shirk once a check rating is achieved. But it would not reject the hypothesis that some firms behave in this manner.

Tables VII presents the probit results for the price and quality ratings. In the column labeled Marginal Probability, the first number represents the marginal probability that firms check rated in the earliest ratings period will repeat their performance in the next rating period. The second number is the marginal probability based on the third vs. second rating period, and so forth. Thus, for an industry that WCC has rated on N occasions, there will be N-1 entries in this column. (The various marginal probabilities were not generated when the initial probit equation showed no significant difference between the probabilities that check rated and uncheck rated firms would be check rated in the next period. Two industries, TV Repair and Tree Experts, do not appear in Table VII because WCC has rated member firms only once. Restaurants and Supermarkets are excluded because WCC did not check rate firms for either price or quality. Thus, the total number of industries in Table VII is 15.)

Table VII shows that firm quality performance is considerably more consistent over timethan price performance. This is not surprising given inevitable measurement errors and other problems in the WCC price index data set. Nevertheless, the price probits reveal a fair amount of consistency in firm ratings. For four industries, the marginal probabilities for all rating periods are significant, and usually highly so. There are mixed results in five industries, with significant marginal probabilities in six periods, six periods with insignificant results, and one period with insufficient observations to permit probit analysis. In one industry, Auto Body Repair, all periods were insignificant. For the remaining four industries, there were either an insufficient number of observations or firms were not check rated for price.

The consistency of firm quality performance is much more impressive. Eleven industries report all marginal probabilities significant in all periods where sample sizes permit probit analysis. One industry, Watch Repair, displayed mixed results. Three industries with small sample sizes report no periods with significant results.

Industry*	Period	Marginal Probability, Price Check	P Value	Marginal Probability, Quality Check	P Value
Auto Body Repair	82-85		>.10	.645	.000
Ruto Body Repair	85-95		>.10	.353	.005
	95-98		>.10	.536	.000
AC & Heating Contractors	77-82		>.10	Insufficient Ol	oservations
C	82-87		>.10	.477	.000
	87-92	.334	.010	.423	.000
	92-97	.450	.000	.350	.020
Appliance Repair	78-85	.485	.025	.484	.000
*	85-91	.440	.047	.423	.009
	91-96	.590	.001	.409	.025
Auto Mechanical Repair	76-81		NA**	.550	.000
_	81-85	.325	.000	.543	.000
	85-88	.330	.000	.492	.000
	88-91	.341	.000	.650	.000
	91-94	.449	.000	.657	.000
	94-97	.367	.000	.555	.000
Carpet Cleaners	87-94		>.10	Insufficient O	oservations
	94-98	.620	.013	.340	.024
Carpet Installers	89-96	Insufficient O	bservations		>.10
Computer Repair	89-94	Insufficient O	bservations	Insufficient O	bservations
	94-98	.576	.045	.420	.080
Drycleaners	90-96	No Price Cheo	ek Variable	.430	.000

TABLE VII Probability That Firms Check Rated For Price or Quality Will Be Check Rated in Subsequent Evaluation Period

Industry	Period	Marginal Probability, Price Check	P Value	Marginal Probability, Quality Check	P Value
Electricians	90-96	Insufficient Ob	oservations	.500	.041
Local Movers	81-92 92-98	Insufficient Ob Insufficient Ob		Insufficient (>.10 Observations
Locksmiths	86-94 94-99	.420	>.10 .070	Insufficient (Observations >.10
Pest Control	77-86 86-93 93-97	 .544	>.10 .10 .009	Insufficient (Insufficient (.681	
Plumbers	77-83 83-89 89-95	.260 .250 .174	.009 .050 .100	.561 .268 .467	.000 .020 .000
Shoe Repair	88-95	.451	.001	.283	.041
Watch Repair	89-91 91-96	Insufficient O No Price Chec			>.10 .002

TABLE VII (Cont.) Probability That Firms Check Rated For Price or Quality Will Be Check Rated in Subsequent Evaluation Period

*Only one ratings period available for TV Repair and Watch Repair. No Price Check or Quality Check variables available for Supermarkets.

**Firms not check rated separately for price in 1976 ratings.

VII. CONCLUSIONS

The results of this Working Paper paint a very mixed picture of the economic performance of the consumer service sector in the Washington D.C. area. Judging strictly from the absolute levels of satisfaction reported by respondents to the <u>Washington Consumer</u> <u>Checkbook</u> questionnaires, consumers appear reasonably happy with the quality of the services they are receiving. Although the data do not allow any rigorous conclusions in this regard, the observed average satisfaction levels indicate that the service sector has avoided anything resembling a "lemons" equilibrium, where only poor quality is provided.

Market efficiency appears much lower, however, when analyzed more formally using the degree of correlation between price and quality as a benchmark. Only a few industries display positive price-quality correlations that are statistically significant and consistent over time. Of even more concern, the plurality of industries report consistently significant *negative* price-quality correlations. This outcome is contrary to any accepted economic theory of markets, and also fails on common sense grounds. There is no reason to believe that the market would perpetually reward the worst performing firms in an industry.

Thus, it is difficult to reject that conclusion that there are data problems afoot. Specifically, respondents to the WCC survey may to some extent be allowing their satisfaction with the price that is charged for a service color their appraisal of the quality of that service. In effect, consumers may be providing "value" rankings rather than absolute quality rankings. This could explain why the WCC quality ratings frequently appear to fall as price increases.

Any such interaction between price and quality ratings cannot, however, explain other results in our study that are inconsistent with certain theories posited in the literature on the economics of information. Irrespective of price, firms that advertise intensively by running large Yellow Pages displays do not perform better in the WCC ratings, and frequently do more poorly than firms that just run simple one-line listings.

In addition, consumers cannot rely on the relative size of a firm as a clue to probable performance. On average, larger firms score lower than smaller firms on quality, and do not differ significantly on price. Nor does affiliation with a national chain or status as a franchised dealership for a manufacturer signal higher quality. Indeed, in the automobile industry, such firms score much lower in consumer satisfaction than do independent service establishments. Finally, in the one industry were data are available, the length of warranty protection offered for a service does not serve as a signal of higher quality.

Our analysis of time series data for firm price and quality performance shows a fairly high level of consistency in the ratings, particularly for quality. Firms that WCC singles out for quality check rating in one rating period display a much higher probability of receiving a check rating in the next period than do firms that initially are not check rated. Thus, in the long run, it does not appear that most firms "milk" any reputation advantage from the WCC ratings by lowering quality.

Overall conclusions are difficult, particularly because of the potentially serious bias introduced by an interaction between price and the WCC quality rankings. It does appear, however, that consumers in the Washington D.C area have a higher probability of reporting a favorable quality assessment to <u>Washington Consumer Checkbook</u>, if they patronize smaller independent service providers and rely on word-of-mouth reputation rather than Yellow Pages displays, or other indirect clues such as firm size, affiliation with a national chain, or status as a franchised dealership.

REFERENCES

1. AKERLOF, GEORGE A., (1970), "The Market for 'Lemons': Quality Uncertainty and the Market Mechanism," <u>Quarterly Journal of Economics</u> 84, 488-500.

2. ARCHIBALD, ROBERT B., CLYDE A HAULMAN, and CARLISLE E MOODY, JR., (1983), "Quality, Price, Advertising, and Published Quality Ratings," <u>Journal of Consumer</u> <u>Research</u> 9, 347-356.

3. BOND, RONALD S., ET AL., (1980), "Effects of Restrictions On Advertising and Commercial Practices In the Professions: The Case of Optometry," Staff Report to the Federal Trade Commission, Washington, D.C.

4. BUTTERS, GERARD R., (1977), "Equilibrium Distribution of Sales and Advertising Prices," <u>Review of Economic Studies</u> 44, 465-491.

5. CAVES, RICHARD E., and DAVID P. GREENE, (1996), "Brands' Quality Levels, Prices, and Advertising Outlays: Empirical Evidence on Signals and Information Costs," <u>International Journal of Industrial Organization</u> 14, 29-52.

6. GERSTNER, EITAN, (1985), "Do Higher Prices Signal Higher Quality?," <u>Journal of</u> <u>Marketing Research</u> 22, 209-215.

7. GEISTFELD, LOREN V., (1988), "The Price Quality Relationship: The Evidence We Have, The Evidence We Need," <u>The Frontier of Research in The Consumer Interest</u>, ACCI: 143-172.

8. IPPOLITO, PAULINE M., (1990), "Bonding and Nonbonding Signals of Product Quality," Journal of Business 63, 41-60.

9. KLEIN, BENJAMIN, and KEITH B. LEFFLER, (1981), "The Role of Market Forces in Assuring Contractual Performance," Journal of Political Economy 89, 615-641.

10. MORRIS, RUBY TURNER, and CLAIRE SEKULSKI BRONSON, (1969), "The Chaos of Competition Indicated by <u>Consumer Reports</u>," <u>Journal of Marketing</u> 33, 26-34.

11. NELSON, PHILLIP, (1974), "Advertising as Information," Journal of Political Economy 81, 729-754.

12. OXENFELDT, ALFRED R., (1950), "Consumer Knowledge: Its Measurement and Extent," <u>Review of Economics and Statistics</u>, 32, 300-314.

13. ROTHCHILD, MICHAEL, (1973), "Models of Market Organization with Imperfect Information: A Survey," Journal of Political Economy 81, 1283-1308.

14. ROTHCHILD, MICHAEL, and JOSEPH STIGLITZ, (1976), "Equilibrium in Competitive Insurance Markets: An Essay on the Economics of Imperfect Information," <u>Quarterly Journal of Economics</u> 90, 629-650.

15. SPENCE, MICHAEL, (1974), "Competitive and Optimal Responses to Signals: An Analysis of Efficiency and Distribution," Journal of Economic Theory 7, 296-332.

16. STIGLER, GEORGE J., (1961), "The Economics of Information," <u>Journal of Political Economy</u> 69, 213-225.

APPENDIX A DETAILED REGRESSION RESULTS

This appendix presents the principal regression results for the most recent ratings period for all industries in the data set. As explained in the main text (see pp. 13-14), the regressions that use the WCC quality ratings as a dependent variable employ a form of logit estimation suitable for use with grouped data. In essence, this procedure unfolds the summary quality rating for a given firm into a series of binary categorical variables, with a zero-one value generated for each respondent rating used in calculating the overall quality score.

Consider, for example, a firm that receives an overall performance rating of 80. For most industries this would mean that 80 percent of the consumers reporting on that firm rated it superior in overall performance. If there were a total of 10 respondents, the logit program would generate a series of dependent variables comprised of eight "one" values and two "zero" values. Each of these dummies would be associated with the corresponding firm-specific independent variables used in the logit regression. Thus, firms with the highest number of individual ratings automatically would be weighted heaviest in the regression estimation procedure. The estimation technique also adjusts standard errors to reflect the expected correlation of error terms among the observations for any given service provider.

The discussion first focuses on regressions that include each of the principal independent variables separately as predictors of the WCC quality score. These bivariate regressions test whether consumers can rely on the attribute in question in isolation as a signal of quality. For the Yellow Pages and firm size variables, regression results are also shown testing the hypothesis that consumers can rely upon the size of a Yellow Pages ad or the size of a firm to signal value, as measured by the WCC quality score divided by the WCC price index.

For those industries with a full complement of independent variables, results are then reported for multivariate regressions that reveal more precisely any independent explanatory power that the various variables might have in predicting firm quality. These results are not, however, directly relevant to the primary signaling hypothesis in question, since such theories do not posit that consumers will consciously or unconsciously control for other factors when viewing a single firm attribute as a possible signal of quality.

Air Conditioning and Heating Contractors (1997)

This data set includes both the Yellow Pages and employment size variables. Regression 1 documents that price and quality are negatively correlated in this industry. In Regression 2, quality and firm size are shown to be negatively correlated at a very high level of significance. Regression 3 shows a positive but insignificant relationship between quality and size of Yellow Pages advertising. Regression 4 indicates that consumers cannot expect a firm with a large

Yellow Pages ad to offer better value in terms of quality per dollar. There is actually a significant negative correlation between these two variables. Regression 5 reveals a negative but insignificant correlation between value (the quality rating adjusted for price) and the size of the firm.

Regression 6 discloses a significant positive relationship between Yellow Pages advertising and quality when firm size and price are also included as independent variables. As is shown in Regression 7, the Yellow Pages variable is positively correlated with price and firm size. Because these two variables are in turn negatively correlated with quality, Yellow Pages functions as a partial proxy for price and size in Regression 3 and loses much of its independent positive correlation with quality.

Regression 1: Dependent Variable = %rating firm superior Independent Variable = WCC price index

Logit estimate	es			Wald	011112(1)	$= 119 \\ = 4.49 \\ = 0.0340$
Log likelihood	d = -4318.1522	1		Pseud	do R2	= 0.0047
		(standard	errors	adjusted fo	or clusterin	g on idno97)
 supq 		Robust Std. Err.	. 2	P> z	[95% Con	f. Interval]
price97 _cons	0098648 2.038521	.0046538 .4890684	-2.1 4.1		0189861 1.079964	

Regression 2: Dependent Variable = %rating firm superior Independent Variable = number of employees

Logit estimate	S			Number Wald cl	of firms hi2(1)	= 13 = 17.7	
Log likelihood	= -4936.9785			Prob > Pseudo		= 0.000 = 0.00	
		(standard e	errors adju	usted for	clusterir	ng on idno97	7)
 supq	Coef.	Robust Std. Err.	Z	P> z	[95% Cor	nf. Interval	L]
employ97 _cons	0082284 1.168227	.0019541 .0950933	-4.21 12.29	0.000 0.000	0120584 .9818473		

Regression 3: Dependent Variable = %rating firm superior Independent Variable = size of Yellow Pages ad

Logit estimates			Number Wald c	of firms = hi2(1) =	115 0.17
Log likelihood = -4624.8634	Ł		Prob > Pseudo		0.6767 0.0002
	(standard	errors	adjusted for	clustering	on idno97)
 supq Coef.	Robust Std. Err.	Z	P> z	[95% Conf	. Interval]
yp97 .0013026 _cons .9231844	.003124 .1079349	0.4 8.5		0048203 .711636	.0074255 1.134733

Regression 4: Dependent Variable = size of Yellow Pages ad Independent Variable = value (quality/price)

Source	SS	df	MS		Number of firms = 102 F(1, 100) = 13.86
Model Residual Total	5610.21758 40478.981 46089.1985	100 404	.21758 .78981 328698		Prob > F = 0.0003 R-squared = 0.1217 Adj R-squared = 0.1129 Root MSE = 20.119
yp97	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
value _cons	-45.44601 57.25293	12.20733 11.50368	-3.723 4.977	0.000 0.000	-69.66501 -21.22702 34.42995 80.07591

Regression 5: Dependent Variable = number of employees Independent Variable = value (quality/price)

Source	SS	SS df MS			Number of firms = 117 F(1, 115) = 1.66	
Model Residual	401.425903 27798.822	1 401. 115 241.	4 2 5 9 0 3 7 2 8 8 8 7		Prob > F R-squared Adj R-squared	$= 0.2001 \\ = 0.0142$
Total	28200.2479	116 243.			Root MSE	= 15.548
employ97	Coef.		 t	P> t	[95% Conf.	Interval]
value _cons	-10.79503 26.18363	8.376937 8.028892	-1.289 3.261	0.200 0.001	-27.38812 10.27994	5.798074 42.08732

Regression 6:	-		number c	-	es, size of	Yellow	
Logit estimate	S				of firms = hi2(3) =	101 14.36	
				Prob >	chi2 =	0.0025	
Log likelihood = -3994.811 Pseudo R2 = 0.0143							
		(standard e	rrors ad	justed for	clustering	on idno97)	
1		Robust					
supq	Coef.	Std. Err.	z	P> z	[95% Conf.	Interval]	
employ97	0141104	.0048582	-2.90	0.004	0236323	0045885	
yp97	.0089862		2.69	0.007	0024337	.0155386	
price97		.004524		0.001	0237323	0059986	
-	2.647565		5.31	0.000	1.669567		
_cons	2.04/505	.4909870	5.31 		1.009507	3.025502	

Regression 7: Dependent Variable = size of Yellow Pages ad Independent Variables = number of employees, WCC price index, %rating firm superior

Source	SS	df	MS		Number of fir F(3, 97)	
Model Residual	13071.5361 32845.053		7.1787608794		Prob > F R-squared Adj R-squared	= 0.0000 = 0.2847
Total	45916.5891	100 459.	165891		Root MSE	= 18.401
yp97	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
employ97 price97 super97 _cons	.4527038 .5658379 .1546826 -61.42468	.1165674 .1158479 .1265388 17.26251	3.884 4.884 1.222 -3.558	0.000 0.000 0.225 0.001	.2213497 .3359119 0964618 -95.68598	.6840579 .7957638 .405827 -27.16337

Auto Body Repair (1998)

The simple two-way regressions for the 1998 Auto Body Repair data set reveal no significant association between price and quality (Regression 1), and highly significant negative correlations between quality and firm size (Regression 2), dealer status (Regression 6), and chain status (Regression 7). (There is only one national chain represented in this data set). Regression 3 reveals no significant relationship between quality and the size of Yellow Pages advertising, and Regression 4 shows that the Yellow Pages variable is not significantly correlated with the value variable (WCC quality score/WCC price index). In Regression 5, there is a highly significant negative correlation between value and size of firm.

With all of the independent variables included in Regression 8, dealer status and chain status continue to display highly significant negative coefficients. Firm size and Yellow Pages advertising lose significance, however, and price remains insignificant. Further investigation revealed that the dealer dummy variable and firm size are positively correlated (r=.42). Thus, when firm size is used as the only regressor, it functions partially as a proxy for dealer status in predicting firm quality. Taken together, the results for the full regression and simple regressions indicate that dealer status is a more powerful predictor of firm quality than is firm size.

In Regression 9, Yellow Pages advertising is removed in order to increase the sample size. (A number of smaller firms rated by WCC could not be located in the Yellow Pages listings.) In the expanded set, firm size achieves a high level of significance, although the z score for dealer status is once again higher than for firm size, and the positive coefficient on the price variable attains significance.

Regression 1: Dependent Variable = %rating firm superior Independent Variable = WCC price index

Logit estimate	S				of firms = hi2(1) =	120 0.36
Log likelihood	= -1935.186	8		Prob > Pseudo	chi2 =	0.0100
		(standard	errors ad	justed for	clustering	on idno98)
 supq	Coef.	Robust Std. Err.	Z	P> z	[95% Conf	. Interval]
price	0065402 1.584347	.0108856 1.106101	-0.60 1.43	0.548	0278756 5835702	.01479533.752264

Regression 2: Dependent Variable = %rating firm superior Independent Variable = number of employees

Logit estimates	3			Number Wald cl Prob >	. ,	= = =	138 15.09 0.0001
Log likelihood		Pseudo		=	0.0155		
		(standard e	errors adju	isted for	clusterir	ng o	n idno98)
 supq 		Robust Std. Err.	Z	P> z	[95% Cor	nf.	Interval]
employ98 _cons	0558852 1.439325	.0143885 .1341409	-3.88 10.73	0.000	0840862 1.176414	_	0276842 1.702237

Regression 3: Dependent Variable = %rating firm superior Independent Variable = size of Yellow Pages ad

Logit estimate	es			Numb	er of firm	s =	95
				Wald	l chi2(1)	=	0.79
				Prob	> chi2	=	0.3730
Log likelihood = -1882.6538 Pseudo R2 = 0.0009							0.0009
		(standard	errors	adjusted f	or cluster	ing	on idno98)
		Robust					
supq	Coef.	Std. Err.	. z	P> z	[95% Co	onf.	Interval]
	+						
ур	0083301	.0093499	-0.8	9 0.373	02665	56	.0099954
_cons	1.04312	.1049185	9.9	4 0.000	.83748	34	1.248756

Regression 4: Dependent Variable = size of Yellow Pages ad Independent Variable = value (quality/price)

Source	SS	df 	MS		Number of firms = 95 F(1, 93) = 1.29
Model Residual Total	130.662894 9404.25158 9534.91447	93 101.1	62894 20985 43526		Prob > F = 0.2586 R-squared = 0.0137 Adj R-squared = 0.0031 Root MSE = 10.056
ур	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
value _cons	-6.909842 12.17246	6.078726 4.41782	-1.137 2.755	0.259 0.007	-18.98099 5.161304 3.399545 20.94537

Regression 5: Dependent Variable = number of employees Independent Variable = value (quality/price)

Source	SS	df	MS		Number of firms = 120 F(1, 118) = 25.21
Model Residual Total	558.897059 2615.85086 3174.74792	118 22.16	897059 582276 785539		F(1, 118) = 25.21 Prob > F = 0.0000 R-squared = 0.1760 Adj R-squared = 0.1691 Root MSE = 4.7083
employ98	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
value _cons	-12.46466 16.65288	2.482447 1.776893	-5.021 9.372	0.000	-17.38058 -7.548738 13.13414 20.17161

Regression 6: Dependent Variable = %rating firm superior Independent Variable = dealer status (dealer = 1)								
Logit estimate	es			Wal	(_)	=		
Log likelihood	d = -2187.062	4			ob > chi2 eudo R2		0.0000 0.0293	
		(standard	errors a	adjusted	for clust	ering o	on idno98)	
supq	•	Robust Std. Err.						
dealer _cons	9553177	.131402	-7.2	7 0.000	0 -1.21	2861	6977746 1.313697	

Regression 7: Dependent Variable = %rating firm superior Independent Variable = chain status (chain = 1)

Logit estimate	S				c of firms hi2(1)	= 138 = 53.28
Log likelihood	= -2242.4496	5		Prob > Pseudo		= 0.0000 = 0.0047
		(standard	errors adj	usted for	clusterin	g on idno98)
 supq +		Robust Std. Err.	Z	P> z	[95% Con	f. Interval]
chain _cons	-1.216412 .9481478	.1666463	-7.30 11.81	0.000	-1.543033 .7907916	8897911 1.105504

Regression 8 Dependent Variable = %rating firm superior Independent Variables = dealer status, wcc price index, size of Yellow Pages ad, number of employees, chain status

Logit estimate Log likelihood		2		LR ch	er of firms = hi2(5) = > chi2 = do R2 =	95 131.31 0.0000 0.0407
supq	Coef.	Std. Err.	Z	P> z	[95% Conf.	Interval]
dealer price yp employ98 chain _cons	-1.085687 .0057761 .0026465 0178601 -1.538303 .7754602	.1257825 .0067361 .006309 .0101569 .3354031 .6555708	-8.63 0.86 0.42 -1.76 -4.59 1.18	0.000 0.391 0.675 0.079 0.000 0.237	-1.332216 0074264 0097188 0377674 -2.195681 5094349	8391573 .0189786 .0150119 .0020471 8809248 2.060355

Regression 9: Dependent Variable = %rating firm superior Independent Variables = dealer status, WCC price index, number of employees, chain status

Logit estimate Log likelihood		LR ch	> chi2 =	120 176.25 0.0000 0.0455		
supq	Coef.	Std. Err.	Z	P> z	[95% Conf	Interval]
dealer price employ98 chain _cons	956223 .010555 0289016 -1.4128 .4006843	.1047231 .0059654 .0080113 .2672092 .5855493	-9.13 1.77 -3.61 -5.29 0.68	0.000 0.077 0.000 0.000 0.494	-1.161477 0011369 0446034 -1.93652 7469713	7509695 .0222469 0131997 8890793 1.54834

Auto Mechanical Repair (1997)

The Automobile Mechanical Repair data set is the largest in the sample, both in terms of the number of firms and the number of possible predictors of firm quality. The results for the full set of independent variables for 1997 have already been discussed in the main text on page 24. The results shown below are for the simple two-way regressions and a regression using size of Yellow Pages advertising as the dependent variable.

Regression 1 reports that price functions as a perverse indicator of quality in this industry. The price coefficient is negative and highly significant. Regression 2 reveals that firm size (as measured by number of employees) is also inversely related to quality. Indeed, all of the remaining signaling variables--size of Yellow Pages ad, dealer status, and chain status-display negative coefficients (Regressions 3, 6, and 7). Regressions (not shown) using dummy variables for various certifications revealed significant negative coefficients for employment of ASIE-certified technicians, and designation as a state safety inspection site, and a positive but insignificant coefficient for certification by the American Automobile Association. Regression 4 reveals a significant negative correlation between size of Yellow Pages ad and value, which is measured by the ratio of the WCC quality score to the WCC price index. In Regression 5, there is a much stronger negative correlation between the value variable and size of firm.

Regression 1:	Dependent Va Independent				or or adequat	e
Logit estimate	25			Wald c	c of firms = chi2(1) = chi2 =	30.05
Log likelihood	a = −6770.2166				R2 =	
		(standard	errors ac	ljusted fo	or clustering	on comno)
 supq 	Coef.					Interval]
pricin97	0159178 3.679106	.0029039	-5.48	0.000	0216093	
Regression 2:	Dependent Va Independent		-	_	—	e
Logit estimate	25			Wald c	c of firms = chi2(1) = chi2 =	55.66
Log likelihood	a = -7610.5535			Pseudo	> chi2 = > R2 =	0.0311
		(standard	errors ac	ljusted fo	or clustering	on comno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Conf.	Interval]
	0429339 2.583182	.0872403		0.000	0542129 2.412194	0316549 2.75417
Regression 3:	Dependent Va	riable = %1	rating fin	rm superio	or or adequat ges advertise	
Logit estimate	25			Wald c	cof firms = chi2(1) = chi2 =	1.11
Log likelihood	d = -4669.489			Pseudo	• chi2 = • R2 =	0.0004
		(standard	errors ac	djusted fo	or clustering	on comno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Conf.	Interval]

		nobube				
supq	Coef.	Std. Err.	Z	₽> z	[95% Conf.	Interval]
1 1	0030682 2.154566		-1.06 25.38		0087671 1.988206	.0026306 2.320926

Regression 4: Dependent Variable = size of Yellow Pages ad Independent Variable = value (quality/price)

Source	SS	df 	MS		Number of firms = 278 F(1, 276) = 3.23
Model Residual + Total	1338.41317 114319.277 115657.69	276 414.	.41317 200277 536786		Prob > F = 0.0733 R-squared = 0.0116 Adj R-squared = 0.0080 Root MSE = 20.352
iotai	119097.09	2// 11/.	550700		ROOT MSE - 20.332
yp97	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
value _cons	-10.3485 19.89839	5.756882 5.525921	-1.798 3.601	0.073	-21.68147 .9844806 9.02008 30.77669

Regression 5: Dependent Variable = number of employees Independent Variable = value (quality/price)

Source	SS	df	MS		Number of firm F(1, 442)	
Model Residual + Total	6026.86996 18524.6993	1 6026 442 41.9	.86996 110844 		Prob > F R-squared Adj R-squared Root MSE	= 0.0000 = 0.2455
1 1	Coef.			P> t	[95% Conf.	Interval]
value _cons		1.384129 1.299164	-11.992 18.519	0.000	-19.31838 21.50629	-13.8778 26.6129

Regression 6: Dependent Variable = %rating firm superior or adequate Independent Variable = dealer status

Logit estimate	S			Wald	per of firm d chi2(1)	=	515 137.69
Log likelihood	= -7531.3521				o > chi2 1do R2	=	0.0000 0.0412
		(standard	errors	adjusted	for cluste	ring	on comno)
 supq	Coef.	Robust Std. Err.	Z	P> z	[95% C	onf.	Interval]
dealer97 _cons	-1.098344 2.591357	.0936021 .0745284	-11.73 34.77		-1.2818 2.4452		9148873 2.73743

Regression 7:	Dependent Va Independent		-	-	ior or ad	lequate	2
Logit estimate	s			Wald	per of fir d chi2(1) p > chi2	=	
Log likelihood	l = -7831.9797				ido R2	=	0.0029
		(standard	errors a	djusted	for clust	ering	on comno)
 supq 	Coef.	Robust Std. Err.	Z	P> z	[95%	Conf.	Interval]
chain97 _cons	6145223 2.021323		-3.79 35.20				2965818 2.133885

Carpet Cleaning (1998)

The only variables in this data set are price and size of Yellow Pages advertisement. Price proves to be positively correlated with quality (Regression 1). Yellow Pages advertising displays an insignificant negative correlation with quality in Regression 2, and with value (quality adjusted for price) in Regression 3.

Regression 1: Dependent Variable = %rating firm superior or adequate Independent Variable = WCC price index

Logit estimates					r of firms = chi2(1) =	33 6.79
					> chi2 =	
Log likelihood = -893.30575 Pseudo R2 = 0.047						
		(standard er	rors adju	sted for	clustering	on comidno)
ļ		Robust				
supq	Coef.	Std. Err.	Z	P> z	[95% Conf	. Interval]
price98 cons	.030564 -1.467002	.0117299 1.046967	2.61	0.009	.0075738 -3.519019	.0535543

Regression 2: Dependent Variable = %rating firm superior or adequate Independent Variable = size of Yellow Pages advertisement

Logit estimates				r of firms = chi2(1) =	= 38 = 0.91
Log likelihood = -1129.89			> chi2		
	(standard er	rors adju	usted for	clustering	on comidno)
supq Coef.	Robust Std. Err.	Z	P> z	[95% Con:	f. Interval]
yp98 0046769 _cons .7946941		-0.96 2.10	0.340 0.036	0142742 .0517356	.0049203 1.537653

Regression 3: Dependent Variable = size of Yellow Pages Ad Independent Variable = value (quality/price)

Source	SS.	df	MS		Number of firms = 27 F(1, 25) = 2.55
Model Residual	1285.4712 12611.6723		5.4712 166893		Prob > F = 0.1230 R-squared = 0.0925 Adj R-squared = 0.0562
Total	13897.1435	26 534.	50552		Root MSE = 22.46
ур98	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
value _cons	-19.90948 34.77875	12.47226 13.80935	-1.596 2.518	0.123 0.019	-45.59657 5.777614 6.337861 63.21964

Carpet Sales and Installation (1996)

WCC respondents were asked to rate carpet firms both for the quality of the sales advice given in the showroom and for the quality of any installation services that were provided. This data set does not have any information on firm size. Regressions 1 and 2, respectively, show a significant positive correlation between price and the quality of sales service, and a positive but insignificant correlation between the size of a firm's Yellow Pages ad and respondent satisfaction with the sales services.

With respect to the quality of installation services, Regression 3 reveals no significant relationship between price and quality. Note, however, that the sample size for installation services is considerably smaller than for sales services (24 vs. 36). Regression 4 shows a positive but insignificant association between the size of Yellow Pages ads and satisfaction with installation (P=.124). Further regression analysis (not shown) detected no significant relationship between

Yellow Pages ad size and price, or for quality adjusted for price, using either the sales service or installation service satisfaction measure.

Regression 1: Dependent Variable = %rating firm adequate or superior for sales services Independent Variable = WCC price index

Logit estimate	es			Number	of firms =	36	
				Wald c	hi2(1) =	2.88	
				Prob >	chi2 =	0.0896	
Log likelihood	Log likelihood = -279.14817 Pseudo R2 = 0.0191						
5							
		(standard	errors ad	ljusted for	clustering	on idno96)	
		·					
		Robust					
supq	Coef.	Std. Err.	z	P> z	[95% Conf	. Interval]	
price96	.0316688	.0186556	1.70	0.090	0048956	.0682331	
cons		1.894259	-0.53	0.597	-4.714365	2.710996	

Regression 2: Dependent Variable = % adequate or superior, sales services Independent Variable = size of Yellow Pages ad

Logit estimates					r of firms =	42
					chi2(1) = > chi2 =	0.46 0.4966
Log likelihood = -314.43372 Pseudo R2 = 0.0021						
		(standard	errors	adjusted fo:	r clustering	on idno96)
ļ		Robust				
supq	Coef.	Std. Err	. z	₽> z	[95% Conf	. Interval]
ур96	.0120109	.0176681	0.6	8 0.497	0226181	.0466398
_cons	2.197738	.2713814	8.1	0 0.000	1.665841	2.729636

Regression 3: Dependent Variable = %adequate or superior, installation Independent Variable = WCC price index

Logit estimate	25			Wald	ber of firms d chi2(1) b > chi2	= 0.04
Log likelihood	d = -249.35669	9			udo R2	= 0.0001
		(standard	errors	adjusted :	for clusterin	ıg on idno96)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Cor	f. Interval]
price96 _cons	.0023001 1.345761	.0113312 1.146281	0.2		0199086 9009087	

Regression 4:	-	ariable = %ad Variable = s	-	-	-	cion	
Logit estimate	S			Wald c	of firms = hi2(1) =	28 2.33	
Log likelihood = -278.00221 Prob > chi2 = 0.1268 Pseudo R2 = 0.0063							
		(standard er	rors adj	justed for	clustering	on idno96)	
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Conf.	. Interval]	
yp96 _cons	.0190109 1.369419		1.53 7.26	0.127 0.000	0053908 .9998473	.0434127 1.73899	

Computer Repair (1998)

This data set contains only two independent variables--price and Yellow Pages ad size. In Regression 1, price is negatively associated with quality. In Regression 2, no significant correlation is found between quality and the size of a firm's Yellow Pages ad. Further analysis (not shown) revealed a direct and significant correlation between the Yellow Pages variable and price, and no correlation between Yellow Pages ad size and value as measured by quality adjusted for price.

Regression 1: Dependent Variable = %rating firm adequate or superior Independent Variable = WCC price index

Logit estimate	S			Wald o	c of firms chi2(1)	=	27 3.10
Log likelihood	l = -420.2377	9		Prob > Pseudo	> chi2 > R2	=	0.0781 0.0152
		(standard erm	rors adju	sted for	clustering	g or	n comidno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Co	nf.	Interval]
price98 _cons	0208608 3.439096	.0118402 1.14026	-1.76 3.02	0.078	0440673	_	.0023455 5.673964

Regression 2:	-	ariable = %ra Variable = s	-	-	-	erior	
Logit estimate	S			Number	r of firm	s =	18
				Wald o	chi2(1)	=	0.79
				Prob :	> chi2	=	0.3749
Log likelihood	= -396.03579	9		Pseudo	5 R2	=	0.0109
		(standard err	ors adju	sted for	clusteri	ng on 	comidno)
		Robust					
supq ++	Coef.	Std. Err.	Z 	P> z	95% C [onf. 	Interval]
yp98	.02605	.0293559	0.89	0.375	03148	65	.0835865
_cons	.9515454	.2741594	3.47	0.001	.41420	29	1.488888

Drycleaners (1996)

The drycleaners data set is one of the largest in the sample, although it does not contain information on firm size. In Regression 1, the price coefficient is positive but just misses significance (P=.102). No significant correlation is found between the size of a firm's Yellow Pages ad and quality in Regression 2. Regression 3 shows that price and Yellow Pages ad size are positively correlated at a high level of significance. There is no relationship between the Yellow Pages variable and value as measured by the WCC quality score divided by the WCC price score (results not shown).

Regression 1: Dependent Variable = %rating firm superior Independent Variable = WCC price index

Logit estimates Log likelihood = -2696.4198					r of firms chi2(1) > chi2	= = =	211 2.68 0.1018 0.0015
Log likelihood = -2696.4198 Pseudo R2 (standard errors adjusted for clus							
 supq 	Coef.	Robust Std. Err.	Z	P> z	[95% Cor	nf.	Interval]
price96 _cons	.0043811 .0386459	.0026777 .2725427	1.64 0.14	0.102 0.887	0008671 495528		.0096293

Regression 2: Dependent Variable = %rating firm superior Independent Variable = size of Yellow Pages ad									
Logit estimate	s			Number	of firms	з =	188		
				Wald c	:hi2(1)	=	2.14		
				Prob >	• chi2	=	0.1434		
Log likelihood	= -2423.923	3		Pseudo	R2	=	0.0013		
		(standard err	ors adju	sted for	clusterin	ng or	n comidno)		
		Robust							
supq	Coef.	Std. Err.	Z	P> z	[95% Co	onf.	Interval]		
96qy	.0521381	.0356334	1.46	0.143	01770	02	.1219782		
_cons	.4273209	.0758077	5.64	0.000	.278740	05	.5759013		

Regression 3: Dependent Variable = size of Yellow Pages ad Independent Variable = WCC price index

Source	SS	df	MS		Number of fir F(1, 186)	
Model Residual	50.4561121 565.497013	1 50.4 186 3.04	4561121 4030652		Prob > F	= 0.0001 = 0.0819
Total	615.953125				Root MSE	= 1.7436
ур96	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
price96 _cons	.023579 -1.571738	.005788 .59892	4.074 -2.624	0.000	.0121605 -2.753288	.0349975 3901887

Electricians (1996)

The data set for electricians is fairly small, but does contain information on firm size as measured by number of employees. Regression 1 reveals a strong negative correlation between price and quality. Further investigation (not shown) detected no correlations between price and any of the other independent variables. Thus, price is not acting as a proxy for, say, firm size. In Regression 2, firm size and quality are not significantly correlated, although there is a significant negative association between quality and size of Yellow Pages ad in Regression 3. Regression 4 shows that the size of a Yellow Pages ad cannot be used as a signal of value. The Yellow Pages variable is negatively correlated with value (as measured by the WCC quality score divided by the WCC price index). In Regression 5, the value variable is also negatively correlated with firm size, but the relationship is not significant.

Regression 1:		ariable = %r; Variable = N			or				
Logit estimate	es			Wald c	c of firms = chi2(1) = chi2 =	10.87			
Log likelihood	d = -414.351	3			R2 =				
					clustering of				
	Coef.	Robust Std. Err.	Z	P> z		Interval]			
price96	0164893 2.993674	.0050023	-3.30	0.001	0262937	006685			
Regression 2: Dependent Variable = %rating firm superior Independent Variable = number of employees									
Logit estimate Log likelihood		5		Wald c Prob >	c of firms = chi2(1) = chi2 = chi2 = chi2 =	1.65 0.1996			
					clustering of				
	Coef.	Robust	Z						
employ96 _cons	0323265 1.596658		-1.28 8.63		0817232 1.234182	.0170701 1.959135			
Regression 3:	Independent	ariable = %r; Variable = ;	-	Cellow Pag	yes ad				
Logit estimate		8		Wald c	c of firms = chi2(1) = chi2 = chi2 = chi2 =	9.05 0.0026			
-			rors adju	sted for	clustering o	n comidno)			
supq	Coef.		Z	P> z	[95% Conf.	Interval]			
yp96 _cons	0224215 1.622553		-3.01 11.33		0370311 1.341832				

Regression 4: Dependent Variable = size of yellow pages ad Independent Variable = value (%superior/WCC price index)

Source	SS	df	MS		Number of firms = 36 F(1, 34) = 5.19
Model Residual Total	1636.75634 10714.695 12351.4514		.75634 .13809 898611		F(1, 34) = 5.19 Prob > F = 0.0291 R-squared = 0.1325 Adj R-squared = 0.1070 Root MSE = 17.752
 ур96	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
value _cons	-24.25008 33.12038	10.64073 9.236143	-2.279 3.586	0.029 0.001	-45.87465 -2.625506 14.35028 51.89048

Regression 5: Dependent Variable = number of employees Independent Variable = value (%superior/ WCC price index)

Source	SS	df	MS	Number of firms = 32 F(1, 30) = 1.20
Model Residual	31.6064178 790.768582	1 31 30 26	.6064178 .3589527	Prob > F = 0.2822 R-squared = 0.0384 Adj R-squared = 0.0064
Total	822.375		.5282258	Root MSE = 5.1341
employ96		Std. Err	. t P> t	[95% Conf. Interval]
value _cons	-3.87234 8.745292	3.536306 3.044999	-1.095 0.282 2.872 0.007	-11.09444 3.349761 2.526574 14.96401

Locksmiths (1999)

The 1999 data set for Locksmiths is very small and lacks a Yellow Pages advertisement size variable. When this WCC rating was published, staff resources were no longer available to collect the relevant Yellow Pages data. In addition, WCC did not provide firm size data for Locksmiths. In Regression 1, the WCC price index is seen to be negatively correlated with the WCC quality index.

Regression 1:	Dependent Va Independent		-	-	r	
Logit estimate	es				of firms =	25
				Prob >	chi2 =	0.0031
Log likelihood = -280.46412						0.0175
		(standard e	errors adj	usted for	clustering	on idno99)
		Robust				
supq	Coef.	Std. Err.	z	P> z	[95% Conf	. Interval]
price99	0459347	0155294	-2.96	0 003	0763717	0154977
cons		1.594268	3.75		2.854279	9.103695
			J./J		2.034279	9.103095

Major Appliance Repair (1996)

This data set includes variables for price, firm size and size of Yellow Pages ad. As is evident in regressions 1-3, all of these variables are strongly and negatively associated with quality. Further analysis (not shown) revealed no correlation between the Yellow Pages variable and value as measured by quality adjusted for price. In Regression 4, the value variable is negatively correlated with firm size (as measured by number of employees). It should be noted that there is a high degree of direct intercorrelation among the independent variables in this data set. This can be seen in Regression 5, where none of the variables displays a significant coefficient when all are included in the same regression predicting quality.

Regression 1: Dependent Variable = %rating firm adequate or superior Independent Variable = WCC price index

Logit estimate Log likelihood				= = =	46 25.84 0.0000 0.0411		
 supg	Coef	standard err Robust Std. Err.	rors adju 	P> z			n comidno)
price96 cons	0473062	.0093066 1.048761	-5.08 6.87	0.000	065546	8	0290657 9.264364

Regression 2:	-	ariable = %r: Variable = :	-	-	-	r
Logit estimate	S			Wald	r of firms = chi2(1) = > chi2 =	
Log likelihood	= -642.45782	2		Pseud	o R2 =	0.0513
			rors adju	usted for	clustering of	n comidno)
1					[95% Conf.	
•					1093703	
					2.306594	

Regression 3: Dependent Variable = %rating firm adequate or superior Independent Variable = size of Yellow Pages ad

Logit estimate	25			Wald	r of firms chi2(1) > chi2	=	35 6.74 0.0095
Log likelihood		Pseud	o R2	=	0.0203		
		(standard er	rors adju	sted for	clustering	g 01	n comidno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Coi	nf.	Interval]
yp96 _cons	0253057 2.66959	.0097507	-2.60 8.32	0.009	044416	-	0061947 3.298127

Regression 4: Dependent Variable = number of employees Independent Variable = value (quality/price)

regress employ96 value

Source	SS		MS		Number of firm F(1, 41)	
Model Residual	111.030719 545.376258	1 111.0			F(1, 41) Prob > F R-squared Adj R-squared	= 0.0061 = 0.1691
Total	656.406977	42 15.62	287375		Root MSE	= 3.6472
employ96		Std. Err.		P> t	[95% Conf.	Interval]
value _cons	-11.75401 14.47038	4.068373 3.853586	-2.889 3.755	0.006 0.001	-19.97026 6.687909	-3.537767 22.25286

Regression 5:		Variable =		s, WCC pr	e or superio ice index, s	
Logit estimate	S			Number	of firms =	23
				Wald c	:hi2(3) =	2.32
				Prob >	chi2 =	0.5093
Log likelihood	= -214.21189	9		Pseudo	R2 =	0.0043
		(standard er: Robust	rors adju 	sted for	clustering o	n comidno)
supq	Coef.	Std. Err.	Z	P> z	[95% Conf.	Interval]
employ96	024605	.0360032	-0.68	0.494	0951699	.0459599
price96	0134052	.0185372	-0.72	0.470	0497375	.0229271
yp96	.008471	.0205447	0.41	0.680	0317959	.048738
_cons	4.198566	1.760559	2.38	0.017	.7479346	7.649198

Movers Local) (1998)

This data set lacks information on firm size, but does include the Yellow Pages ad size variable. In addition, a dummy variable is used to identify firms that are agents for national moving company chains. This permits a test of the hypothesis that consumers can rely on a firm's status as an agent to signal higher quality. Two price variables are used in the regressions that follow. The first is the usual WCC price index, which in this sample was only available for 22 firms. The second is a firm's hourly wage rate for a crew of 3 during peak moving season, which was available for 26 firms.

Regressions 1 and 2 find no significant correlation between price and quality using either price measure. In Regression 3, Yellow Page ad size is not associated with quality. Further analysis (not shown) revealed that the Yellow Pages variable was not related to value as measured by the WCC quality score divided by the WCC price index or the hourly wage rate variable

Finally, Regression 4 shows that agents for national chains did not receive higher ratings than independent firms. The coefficient for the agent dummy is negative, although it does not achieve significance.

Regression 1:	Dependent Va Independent	ariable = %ra Variable = W			or			
Logit estimate	es			Wald o	r of firms = chi2(1) = > chi2 =	0.00		
Log likelihood	d = -338.9951	2			o R2 =			
					clustering o			
supq	Coef.	Robust Std. Err.	Z		[95% Conf.			
	.0001009 1.06614	.0163045 1.616056	0.01 0.66	0.995 0.509	0318552 -2.101272	.0320571 4.233553		
Regression 2: Dependent Variable = %rating firm superior Independent Variable = hourly wage rate, crew of 3								
Logit estimate		1		Wald o Prob :	r of firms = chi2(1) = > chi2 = o R2 =	1.53 0.2163		
			ors adiu		clustering o			
					-			
supq	1	Debugt						
hour398	 Coef.	Robust Std. Err. .0161001	z 1.24	P> z 0.216	[95% Conf. 	Interval] .0514638		
hour398	Coef. +	Robust Std. Err. .0161001 1.643322	z 1.24 -0.54	P> z 0.216 0.588	[95% Conf. 0116475 -4.110533	Interval] .0514638		
hour398 _cons	Coef. +	Robust Std. Err. .0161001 1.643322 ariable = %ra	z 1.24 -0.54	P> z 0.216 0.588 cm superio fellow Pag Number Wald o	[95% Conf. 0116475 -4.110533 	Interval] .0514638 2.33117 		
hour398 _cons	Coef. 0199081 8896816 	Robust Std. Err. .0161001 1.643322 	z 1.24 -0.54	P> z 0.216 0.588 cm superio fellow Pag Number Wald o	[95% Conf. 0116475 -4.110533 	Interval] .0514638 2.33117 		
hour398 _cons 	Coef. 0199081 8896816 	Robust Std. Err. .0161001 1.643322 	z 1.24 -0.54	P> z 0.216 0.588 cm superio Yellow Pag Number Wald o Prob 3 Pseudo	[95% Conf. 0116475 -4.110533 	Interval] .0514638 2.33117 		
hour398 _cons 	Coef. 0199081 8896816 	Robust Std. Err. .0161001 1.643322 	z 1.24 -0.54	P> z 0.216 0.588 cm superio Yellow Pag Number Wald o Prob 3 Pseudo	[95% Conf. 0116475 -4.110533 	Interval] .0514638 2.33117 		

_ _

Regression 4:	-	ariable = %ra Variable = a	-	-				
Logit estimate	es			Number	c of fir	ms =	30	
-				Wald o	chi2(1)	=	0.37	
				Prob >	> chi2	=	0.5431	
Log likelihood	d = -425.221	2		Pseudo	D R2	=	0.0024	
	(standard errors adjusted for clustering on comidno)							
		Robust						
supq	Coef.	Std. Err.	Z	₽> z	[95%	Conf.	Interval]	
agent98	2566645	.42205	-0.61	0.543	-1.083	8867	.5705383	
_cons	1.09089	.1886116	5.78	0.000	.7212	2183	1.460562	

Pest Control Firms (1997)

The Pest Control data set contains information on firm size and Yellow Pages ad size, as well as two other variables that permit testing of signaling hypotheses. The first is a dummy variable identifying firms that are part of national chains. The second is a variable representing the length of the warranty period that firms offer for pest control work. This variable is used to test the hypothesis that consumers can rely on warranty length as a signal of the effectiveness of the initial treatment. WCC provides two price measures for pest control firms. One is the charge for a termite inspection with written report. The other is more complex and not easily adapted for use as a dependent variable.¹ All price regressions reported below employ the termite inspection charge.

Regression 1 shows a strong positive correlation between consumer satisfaction and price. In Regression 2, however, a similarly strong negative correlation surfaces between satisfaction and firm size and, in Regression 3, between satisfaction and size of Yellow Pages ad. In Regression 4, there is a significant negative relationship between the Yellow Pages variable and value (quality divided by price). The value variable also displays a significant negative coefficient when firm size is used as the dependent variable in Regression 5. Regression 6 reveals a strong negative

¹ This three-part measure consists of (1) a firm's estimated charge for the first treatment of a sample house, (2) the length of period during which the firm will perform a free follow-up treatment, and (3) the estimated charge for treatment after the free followup period. Various alternative specifications were constructed by compiling a composite index that weighted the initial treatment charge by the length of free followup period, and combined this with the charge after the followup period. All such composite measures were either negatively correlated with quality or uncorrelated. In any event, it is not clear which, if any, of the tested specifications is most appropriate, particularly since it is impossible to estimate expected total costs without firmspecific knowledge of the probability that followup treatments will be needed.

correlation between satisfaction and firm status as a member of a national chain. For the final signaling hypothesis, Regression 7 indicates that consumers cannot rely on the length of the warranty period for the initial treatment as a predictor of satisfaction. There is a significant negative correlation between these two variables.

All of the signaling variables are included as predictors of satisfaction in Regression 8. With the exception of size of Yellow Pages ad, all of the variables are significant and retain the same sign as displayed in the simple two way regression. Further analysis showed that the Yellow Pages variable is highly correlated with chain status (r=.52), and loses its explanatory power when both variables are used as predictors.

Regression 1: Dependent Variable = %rating firm superior Independent Variable = termite inspection charge

Logit estimate:	S				r of firms chi2(1)	=	51 16.96
Log likelihood = -1480.7173					> chi2	=	0.0000 0.0173
		(standard er	rors adju	isted for	clusterin	.g o:	n comidno)
 supq	Coef.	Robust Std. Err.	Z	P> z	[95% Co	nf.	Interval]
chgter97 _cons	.0235317 8820023	.0057148 .3129227	4.12 -2.82	0.000 0.005	.012330 -1.49531	-	.0347325 2686851

Regression 2: Dependent Variable = %rating firm superior Independent Variable = number of employees

Logit estimates	3			Wald	r of firms chi2(1)	=	55 17.88
Log likelihood	Prob Pseud	> chi2 .o R2	=	0.0000 0.0227			
		(standard er	rors adju	usted for	clusterin	g o:	n comidno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Co:	nf.	Interval]
empls97 _cons	07324 .8631196	.0173194 .2232701	-4.23 3.87	0.000	107185 .425518	-	0392946 1.300721

Regression 3: Dependent Variable = %rating firm superior Independent Variable = size of Yellow Pages ad

Logit estimate	28	Wald	r of firms chi2(1) > chi2	= 51 = 12.23 = 0.0005		
Log likelihood	d = -1512.411	Pseud	lo R2	= 0.0219		
		(standard er	rors adjı	usted for	clustering	g on comidno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Cor	nf. Interval]
yp97 _cons	0125807 .792496	.0035981	-3.50 3.86	0.000	0196328	

Regression 4: Dependent Variable = size of Yellow Pages ad Independent Variable = value (quality/price)

Source	SS	df 	MS		Number of firms = 48 F(1, 46) = 8.75
Model Residual + Total	5469.12685 28742.3614 34211.4883	46 624.8	.12685 333944 904006		Prob > F = 0.0049 R-squared = 0.1599 Adj R-squared = 0.1416 Root MSE = 24.997
yp97		Std. Err.		P> t	[95% Conf. Interval]
value _cons		7.632881 10.37279	-2.959 5.291	0.005	-37.94634 -7.217952 34.00205 75.76072

Regression 5: Dependent Variable = number of employees Independent Variable = value (quality/price)

Source	SS	df	MS		Number of firms = 51 F(1, 49) = 4.45
Model Residual + Total	66.4344209 731.920183 798.354604	1 66.4 49 14.9	344209 371466 		F(1, 49) = 4.45 Prob > F = 0.0401 R-squared = 0.0832 Adj R-squared = 0.0645 Root MSE = 3.8649
÷ 1	Coef.		t	P> t	[95% Conf. Interval]
value _cons	-2.453904 8.672254	1.163576 1.55833	-2.109 5.565	0.040	-4.7921981156102 5.540673 11.80383

Regression 6:	—	ariable = %r Variable =	-	-				
Logit estimate	S			Number Wald o Prob 2	r of firms = chi2(1) = > chi2 =	57 39.32 0.0000		
Log likelihood	= -1607.7795	5		Pseudo	o R2 =	0.0491		
	(rors adju		clustering on	n comidno)		
		Robust						
	Coef.			P> z	[95% Conf.	Interval]		
					-1.415228			
_cons	.721779	.1444339	5.00	0.000	.4386938	1.004864		
Regression 7: Dependent Variable = %superior, doing work properly Independent Variable = length of warranty for initial treatment								
Logit estimate	S			Number Wald d	r of firms = chi2(1) =	42 3.97		
		_		Prob >	> chi2 =	0.0462		
Log likelihood	l = -518.83793	3		Pseudo	D R2 =	0.0068		
	((standard er	rors adju	sted for	clustering of	n comidno)		
	Coef.			P> z	[95% Conf.	Interval]		
period97				0.046	0030662	0000261		
_cons	2.207526	.1784935	12.37	0.000	1.857685	2.557366		
Regression 8:	Independent	Variables =	price in	ndex, firm chain st	n size, size (tatus			
Logit estimate	S			Wald d	r of firms = chi2(4) = > chi2 =			
Log likelihood	= -1367.2647	7		Pseudo		0.0705		
	((standard er	rors adju	sted for	clustering o	n comidno)		
 supq	Coef.	Robust Std. Err.	Z	P> z	[95% Conf.	Interval]		
+ chgter97	.0123134	.0065977	1.87	0.062	0006179	.0252447		
empls97	057151	.0220553	-2.59	0.010	1003785			
yp97 chain97		.0034326 .2715573	0.46 -3.72	0.646 0.000	0051493 -1.541252			
_cons	.453565	.4369342	1.04	0.299	4028103	1.30994		

Plumbers (1995)

The 1995 data set for Plumbers contains variables for size of firm and Yellow Pages ad. The dependent variable, however, is the less sensitive overall satisfaction measure that combines superior and adequate ratings. A dummy variable was also constructed to identify members of a national chain (of which there was only one.) Analysis (not shown) revealed that this variable was unrelated to consumer satisfaction or to any of the other independent variables.

Regression 1 shows a strong negative correlation between price and quality. Regression 2 reveals a similarly strong negative relationship between quality and firm size. In Regression 3, there is an even more systematic negative relationship between size of Yellow Pages ad and quality. Regression 4 shows that Yellow Pages ad size cannot be used to identify firms providing particularly good value in terms of price-adjusted quality. This reflects a positive correlation between the Yellow Pages variable and price (r=.390). Value is also negatively correlated with firm size, as shown in Regression 5. Finally, Regression 6 reveals that the Yellow Pages variable demonstrates the greatest predictive power in a regression that employs all of the independent variables. This suggests that Yellow Pages ad size is not merely functioning as a proxy for price or firm size in this data set.

Regression 1: Dependent Variable = %rating firm superior or adequate Independent Variable = WCC price index

Logit estimate:	Logit estimates				r of firms =	134
					chi2(1) = > chi2 =	7.37 0.0066
Log likelihood = -1657.9036				Pseud	o R2 =	0.0083
		(standard er:	rors adju	usted for	clustering (on comidno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Conf	. Interval]
price95	0229292	.0084459	-2.71	0.007	0394828	0063755
_cons	4.895038	.8762987	5.59	0.000	3.177524	6.612551

Regression 2: Dependent Variable = %rating firm superior or adequate Independent Variable = number of employees

Logit estimates	5				r of firms : chi2(1)	
Log likelihood = -1758.2104				Prob Pseud	> chi2 o R2	= 0.0125 = 0.0078
		(standard err	ors adju	sted for	clustering	on comidno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Con:	f. Interval]
employ95 _cons	0140669 2.617006	.0056328 .1228796	-2.50 21.30	0.013	0251069 2.376167	0030269 2.857846

Regression 3: Dependent Variable = %rating firm superior or adequate Independent Variable = size of Yellow Pages ad

Logit estimates	5			Wald	r of firms chi2(1)	=	117 13.19
Log likelihood	Prob Pseud	> chi2 o R2	=	0.0003 0.0391			
		(standard er	rors adju	isted for	clustering	g or	n comidno)
 supq	Coef.	Robust Std. Err.	Z	P> z	[95% Cor	nf.	Interval]
yp95 _cons	0187058 2.882208	.0051502	-3.63 22.99	0.000 0.000	0288	-	0086117 3.127879

Regression 4: Dependent Variable = size of Yellow Pages ad Independent Variable = Value (quality/price)

Source	SS	df	MS		Number of firms = 113 F(1, 111) = 35.18
Model Residual Total		1 1663 111 472	1.9631 .72244 		F(1, 111) = 35.18 Prob > F = 0.0000 R-squared = 0.2407 Adj R-squared = 0.2338 Root MSE = 21.742
 ур95	Coef.	Std. Err.		P> t	[95% Conf. Interval]
value _cons	-84.0762 93.41068	14.17439 12.97204	-5.932 7.201	0.000	-112.1637 -55.9887 67.70572 119.1156

Source	SS	df	MS		Number of firms = 133 F(1, 131) = 11.45
Model Residual		1 2268 131 198	.99974 .13534		F(1, 131) = 11.45 Prob > F = 0.0009 R-squared = 0.0804 Adj R-squared = 0.0734
Total	28224.7293				Root MSE = 14.076
employ95		Std. Err.	t	P> t	[95% Conf. Interval]
value _cons	-28.16574	8.323095 7.774728	-3.384 4.358	0.001 0.000	-44.6308 -11.70067 18.50579 49.26632

Regression 5: Dependent Variable = number of employees Independent Variable = value (quality/price)

Regression 6: Dependent Variable = %rating firm superior or adequate Independent Variable = WCC price index, size of Yellow Pages ad, number of employees

Logit estimate		Wald	er of firms chi2(3) > chi2 lo R2	s = 112 = 25.11 = 0.0000 = 0.0267		
		(standard er	rrors adju	sted for	clusterin	ng on comidno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Co	onf. Interval]
price95 yp95 employ95 _cons	0129418 0136778 0037818 4.179633	.0094567 .0045599 .0038631 .9420912	-1.37 -3.00 -0.98 4.44	0.171 0.003 0.328 0.000	031476 02261 011353 2.33316	0047407 .0037897

Restaurants (1998)

This very large data set contains consumer evaluations of 672 restaurants in the Washington D.C. area. Respondents were asked to base their ratings on a scale of 0 through 100, and to rate a restaurant on quality of food, quality of service, and on value. Thus, for this one industry, the quality ratings may not be confounded to any extent with the price ratings, since consumers were specifically asked to provide a separate rating for price-adjusted quality. WCC assigned restaurants to one of five price categories. Although the quality variable is truncated at both ends, the ratings were fairly evenly distributed and did not bunch at the upper truncation point (100). For these reasons, the regressions were run in ordinary least squares format.

Price and food quality are very highly correlated in Regression 1. Note that the coefficients for the four dummy variables (with category 1 the eliminated variable) increase monotonically, showing that quality increases with each increment in the price grouping. The same pattern emerges for quality of service in Regression 2, with even higher significance levels for all of the

coefficients. WCC did not collect firm size data for restaurants, and Yellow Pages ad size data were not collected due to resource constraints and the perceived low probability that such a variable would be correlated with quality in this industry.

Regression 1: Dependent Variable = food quality (0-100) Independent Variable = price category

-	n with robust clusters (idr	Number of fir F(4, 671) Prob > F R-squared Root MSE				
food98	Coef.	Robust Std. Err.	t	P> t	[95% Conf.	Interval]
p298 p398 p498 p598 cons	2.717179 4.282356 7.125349 10.91182 71.77	.6798711 .7127837 .8785988 1.548333 .6033904	3.997 6.008 8.110 7.047 118.945	0.000 0.000 0.000 0.000 0.000 0.000	1.382249 2.882801 5.400215 7.871658 70.58524	4.05211 5.681911 8.850483 13.95198 72.95476

Regression 2: Dependent Variable = service quality (0-100) Independent Variable = price category

-	with robust	Number of fir F(4, 671) Prob > F R-squared Root MSE				
servic98	Coef.	Robust Std. Err.	t	P> t	[95% Conf.	Interval]
p298 p398 p498 p598 cons	3.40044 5.117487 8.467907 11.02182 68.16	.5966455 .6263974 .8089955 1.557838 .513327	5.699 8.170 10.467 7.075 132.781	0.000 0.000 0.000 0.000 0.000 0.000	2.228923 3.887552 6.87944 7.962995 67.15208	4.571956 6.347422 10.05637 14.08064 69.16792

Shoe Repair (1995)

The Shoe Repair data set for 1995 contains price and quality information for 95 firms. Although WCC did not collect information on firm size, the data set does contain a Yellow Pages ad size variable. Regression 1 shows a negative but not quite significant negative correlation between price and quality. In Regression 2, the size of a firm's Yellow Pages ad is positively associated with quality. Further analysis (not shown) revealed no association between the Yellow Pages variable and value as measured by the price-adjusted WCC quality rating.

Regression 1: Dependent Variable = %percent rating firm superior Independent Variable = WCC price index

Logit estimates					r of firms = chi2(1) =	95 2.28
Log likelihood = -		Prob Pseud	> chi2 = o R2 =	0.1310 0.0019		
		(standard er	rors adju	sted for	clustering c	on comidno)
 supq	Coef.	Robust Std. Err.	Z	P> z	[95% Conf.	Interval]
F = = 0 + 0 + 0	0098966	.0065534 .6473709	-1.51 3.95	0.131 0.000	0227411 1.286469	.002948 3.824116

Regression 2: Dependent Variable = %rating firm superior Independent Variable = size of Yellow Pages ad

Logit estimate	S				r of firms = chi2(1) =	77 8.83
Prob > chi2 = 0.0 Log likelihood = -1135.0913 Pseudo R2 = 0.0						
		(standard er	rors adju	sted for	clustering o	on comidno)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% Conf	. Interval]
yp95 _cons	.1727471 1.401833	.0581234 .1263968	2.97 11.09	0.003	.0588274 1.1541	.2866668 1.649566

Supermarkets (2001)

The Supermarkets data set for 2001 includes only 8 firms, but over 10,000 respondents provided ratings based on experiences at dozens of individual stores. Regression 1 shows a very strong positive correlation between price and quality in this industry.

Regression 1: Dependent Variable = % rating store superior Independent Variable = WCC price index

Logit estimates	5			Number	of chains =	8
				Wald ch	i2(1) =	116.89
				Prob >	chi2 =	0.0000
Log likelihood = -6870.3591 Pseudo R2 = 0.03						
		(standard	errors	adjusted fo	r clustering	g on idno)
		Robust				
supq	Coef.	Std. Err.	Z	P> z	[95% Conf.	Interval]
+-						
price	.0413359	.0038233	10.81	0.000	.0338424	.0488294
_cons	-4.419366	.4628962	-9.55	0.000	-5.326626	-3.512106

TV Repair (1989)

The 1989 ratings for TV Repair were the last to include a price index for actual repairs. The 1991 ratings provide information only on a firm's charge for a repair estimate. Neither firm size nor contemporaneous Yellow Pages ad size data could be obtained for the firms in the1989 data set. Regression 1 shows a negative but insignificant correlation between price and the percentage of respondents rating a firm superior or adequate in overall quality.

Regression 1: Dependent Variable = %rating firm superior or adequate

Logit estimates				er of firms =	36
Log likelihood = -372.66854	Ł			chi2(1) = > chi2 = lo R2 =	1.04 0.3072 0.0055
	(standard	errors	adjusted	for clusteri	ng on idno)
supq Coef.	Robust Std. Err.	Z	P> z	[95% Conf	. Interval]
price 0089139 _cons 3.061545	.0087299 .9877203	-1.02 3.10	0.307 0.002	0260241 1.125649	.0081964 4.997442

Tree Experts (1999)

The 1999 data set for Tree Experts is limited to information on price and quality. WCC could not provide firm size data, and resources were no longer available to collect Yellow Pages ad size data. Regression 1 shows no significant correlation between price and quality.

Regression 1: Dependent Variable = %rating firm superior Independent Variable = WCC price index

Logit estimates				Number Wald c Prob >	. ,	29 0.70 0.4029`
Log likelihood =	-505.6894	ł		Pseudo	R2 =	0.0031
		(standard	errors a	djusted for	clustering	on idno99)
 supq 	Coef.	Robust Std. Err.	Z	P> z	[95% Conf.	Interval]
price99 _cons	.0069715 .6667981	.008334 .8065977	0.84	0.403 0.408	0093627 9141043	.0233058 2.247701

Watch Repair (1996)

The 1996 data set for Watch Repair contains information on Yellow Pages ad size, but not firm size. In Regression 1, The WCC quality ratings and the WCC price index are uncorrelated. Similarly, in Regression 2, there is no significant relationship between quality and size of Yellow Pages ad. Further analysis (not shown) failed to find any significant correlation between the Yellow Pages variable and price or value (quality divided by price).

Regression 1: Dependent Variable = %rating firm superior Independent Variable = WCC price index

Logit estimates				c of firms =	41
				chi2(1) =	0.12
			Prob :	> chi2 =	0.7321
Log likelihood = -622.39017			Pseudo	o R2 =	0.0002
	(standard	errors ad	justed for	clustering	on idno96)
	Robust				
supg Coef.	Std. Err.	Z	P> z	[95% Conf	. Interval]
					. incervarj
price96 0018618	.0054376	-0.34	0.732	0125193	.0087958
cons 1.371141	.5230825	2.62	0.009	.3459186	2.396364

Regression 2: Dependent Variable = %rating firm superior Independent Variable = size of Yellow Pages ad							
Logit estimates				Number of obs		=	37
				Wald o	chi2(1)	=	0.01
				Prob :	> chi2	=	0.9065
Log likelihood = -707.40833				Pseudo	5 R2	=	0.0000
		(standard e	rrors adj 	usted for	r cluster	ing (on idno96)
supq	Coef.	Robust Std. Err.	Z	P> z	[95% C	onf.	Interval]
+ ур96	0093096	.0792308		0.906	16459	 91	.1459799
_cons	1.191147	.1316236	9.05	0.000	.9331	69	1.449124